



UNIVERSITAT POLITÈCNICA DE CATALUNYA  
BARCELONATECH

Escola Superior d'Enginyeries Industrial,  
Aeroespacial i Audiovisual de Terrassa

## Nonlinear time series analysis

Master degree in Industrial Engineering

Master degree in Aeronautic Engineering

Course 2020-2021

# Bivariate and multivariate analysis

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## ■ Introduction

- Historical development: from dynamical systems to complex systems

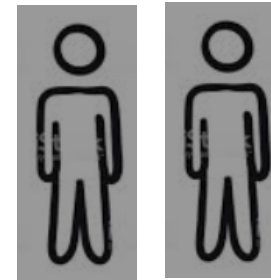
## ■ Univariate analysis

- Methods to extract information from a time series.
- Applications.



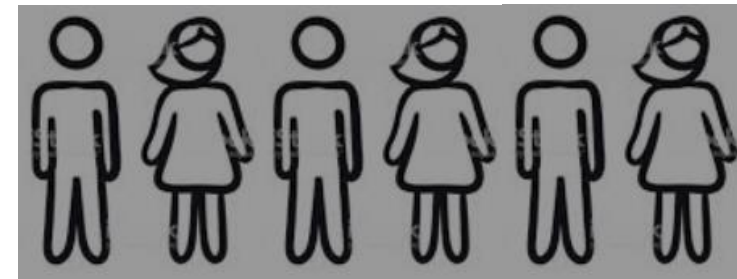
## ■ Bivariate analysis

- Correlation, directionality and causality.
- Applications.



## ■ Multivariate analysis

- Many time series: complex networks.
- Network characterization and analysis.



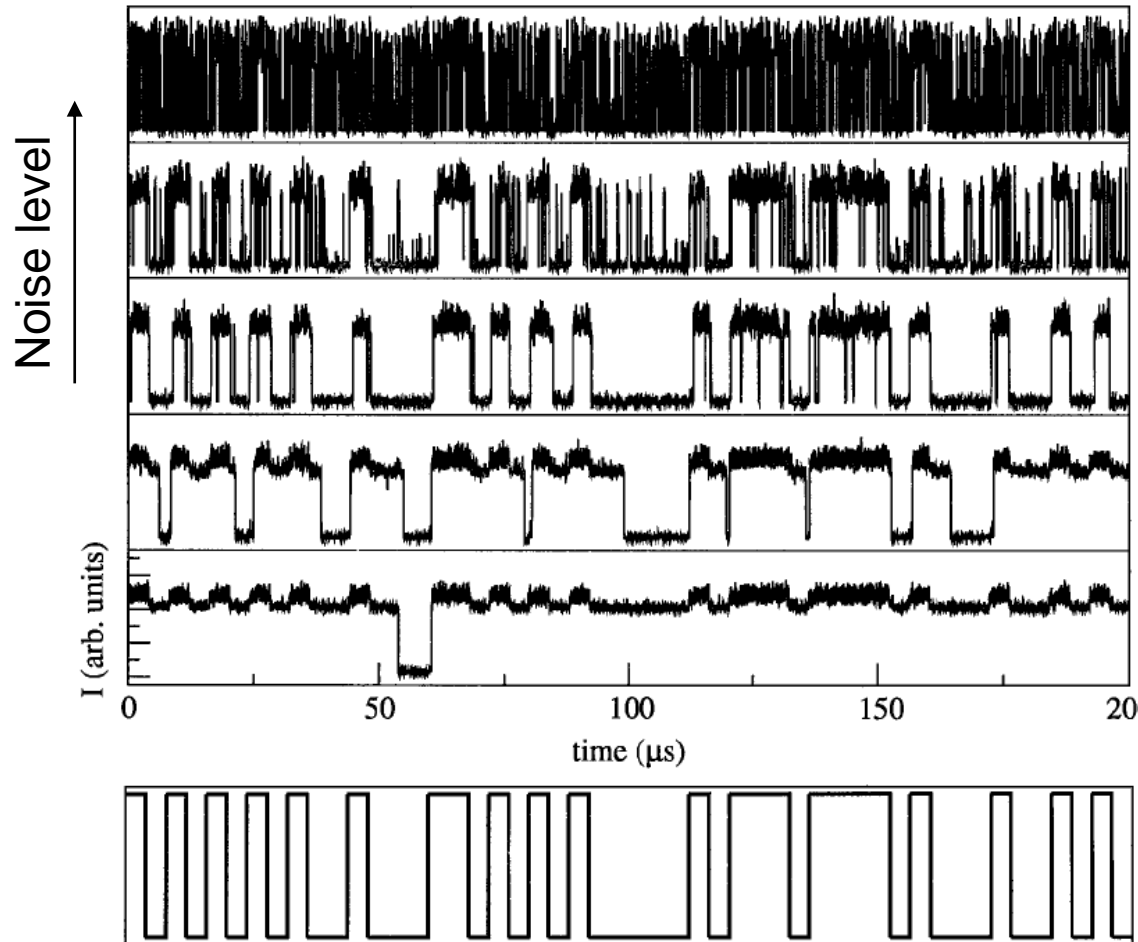
# Cross-correlation of two time series $X$ and $Y$ of length $N$

$$C_{xy}(\tau) = \frac{1}{N - \tau} \sum_{k=1}^{N-\tau} x(k + \tau)y(k)$$

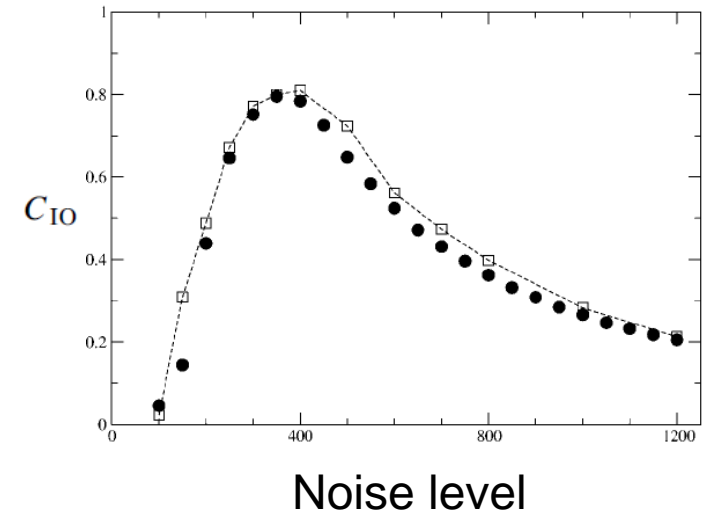
the two time series are normalized to zero-mean  $\mu=0$  and unit variance,  $\sigma=1$

- $-1 \leq C_{X,Y} \leq 1$
- $C_{X,Y} = C_{Y,X}$
- The maximum of  $C_{X,Y}(\tau)$  indicates the **lag** that renders the time series  $X$  and  $Y$  best aligned.
- Pearson coefficient:  $\rho = C_{X,Y}(0)$

# Example: response of a bistable system to an aperiodic signal (stochastic resonance)

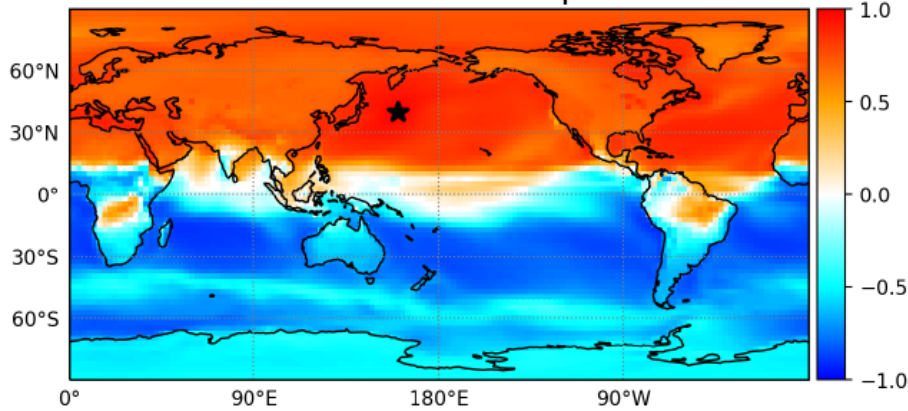


Cross-correlation between input and output signal.

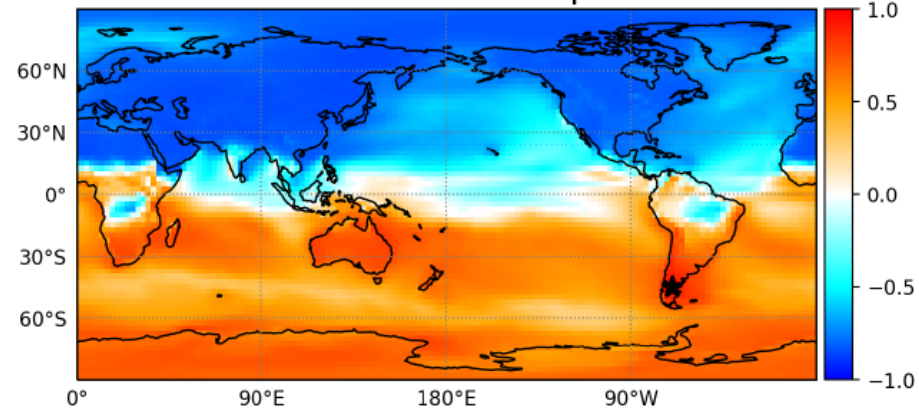


# Example: cross-correlation of cosine of Hilbert phase of SAT at a reference point (\*), and all other regions

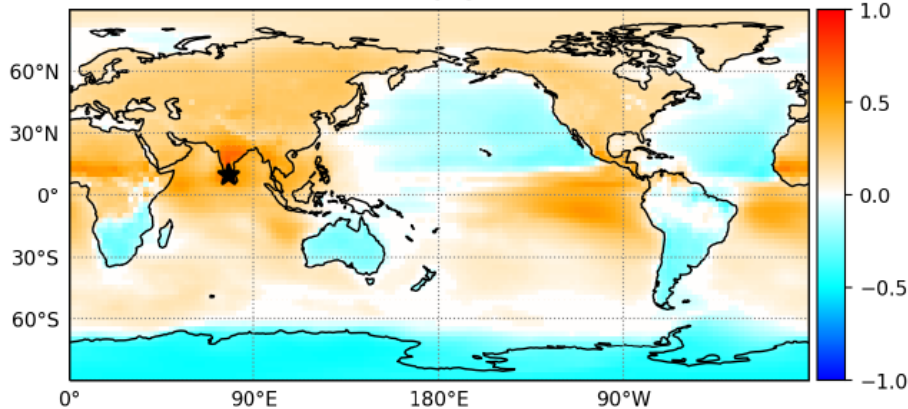
Northern extratropics



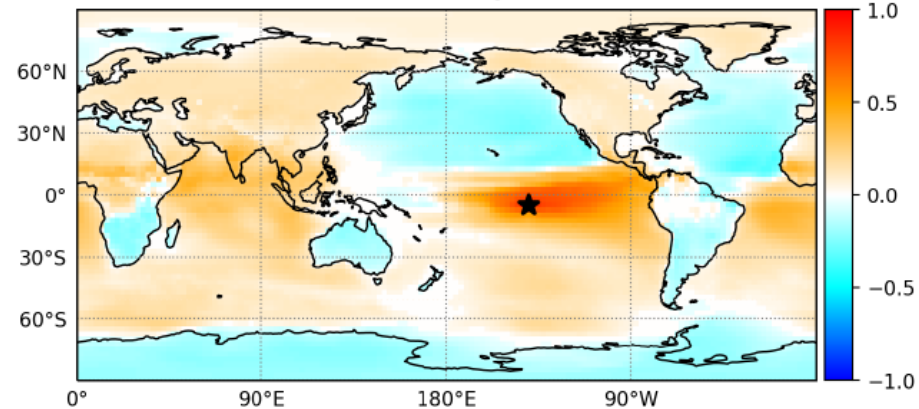
Southern extratropics



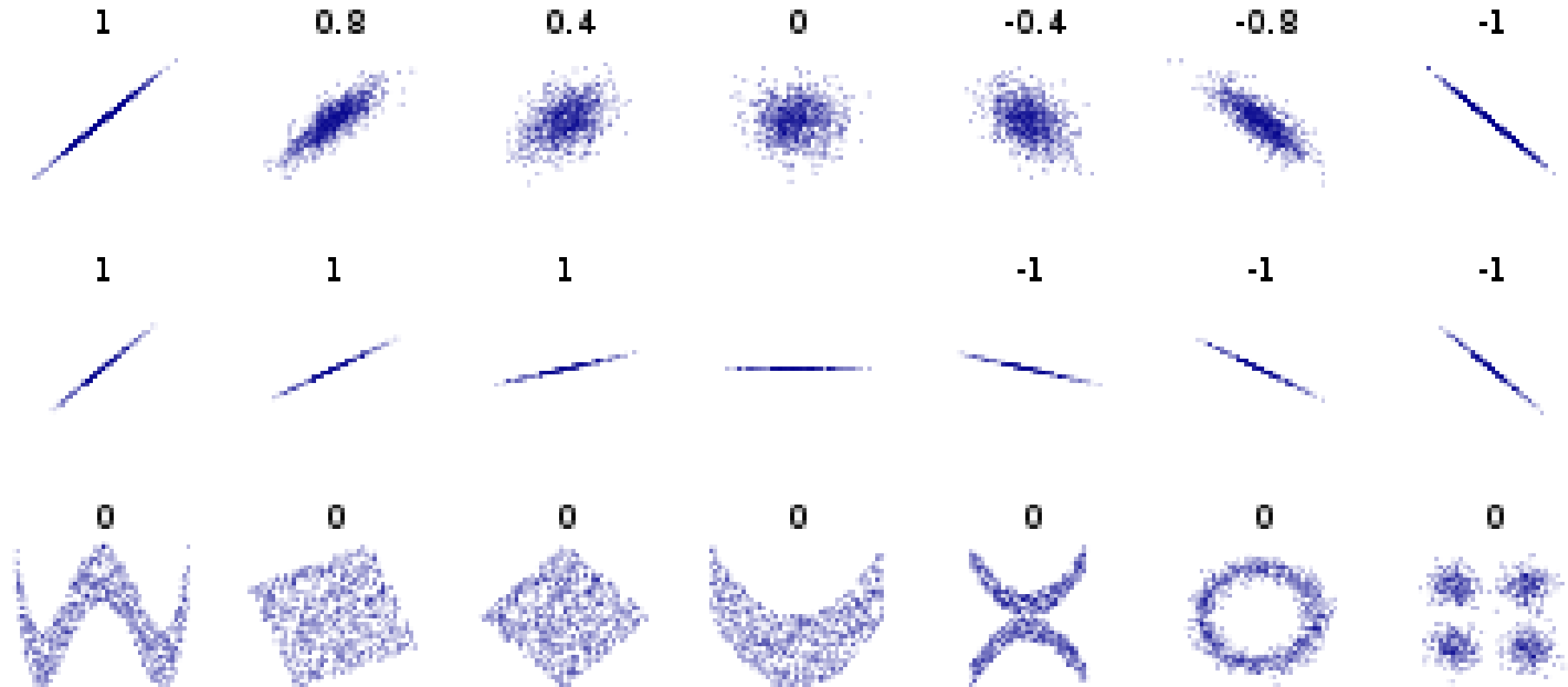
India



El Niño

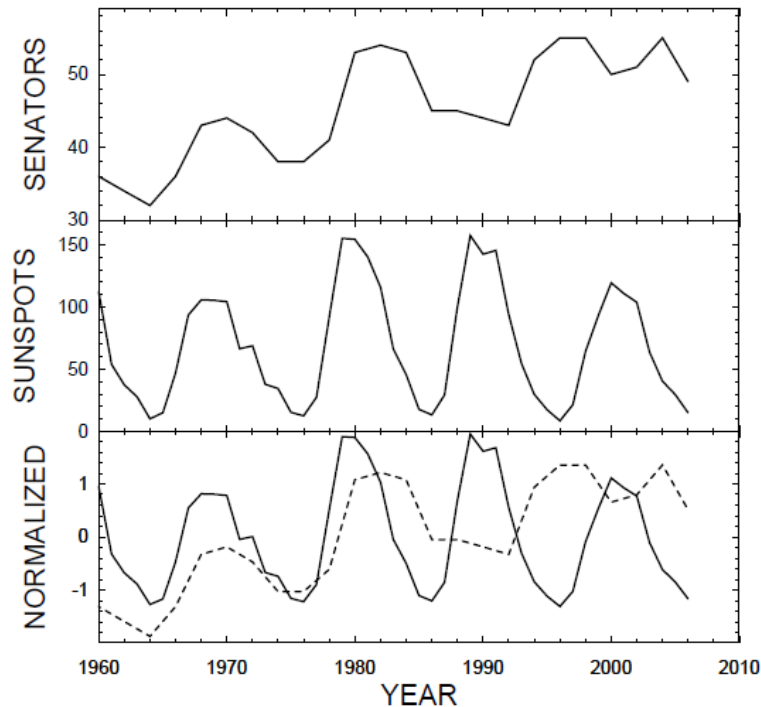


# Cross-correlation analysis detects linear relationships only



# Correlation is NOT causality

An illustrative example: the number of sunspots and the number of the Republicans in the U.S. Senate in the years 1960-2006.



Interval 1960 to 1986 (biannual sampling, 14 points):

**C=0.52** Is this significant?

# Surrogate test

The significance of a correlation value is usually checked by calculating the cross-correlation from an ensemble of signals (**surrogates**) with the same **autocorrelation** than the original time series but completely independent from each other.

<http://tylervigen.com/spurious-correlations>

G. Lancaster et al, “*Surrogate data for hypothesis testing of physical systems*”, Physics Reports 748 (2018) 1–60.



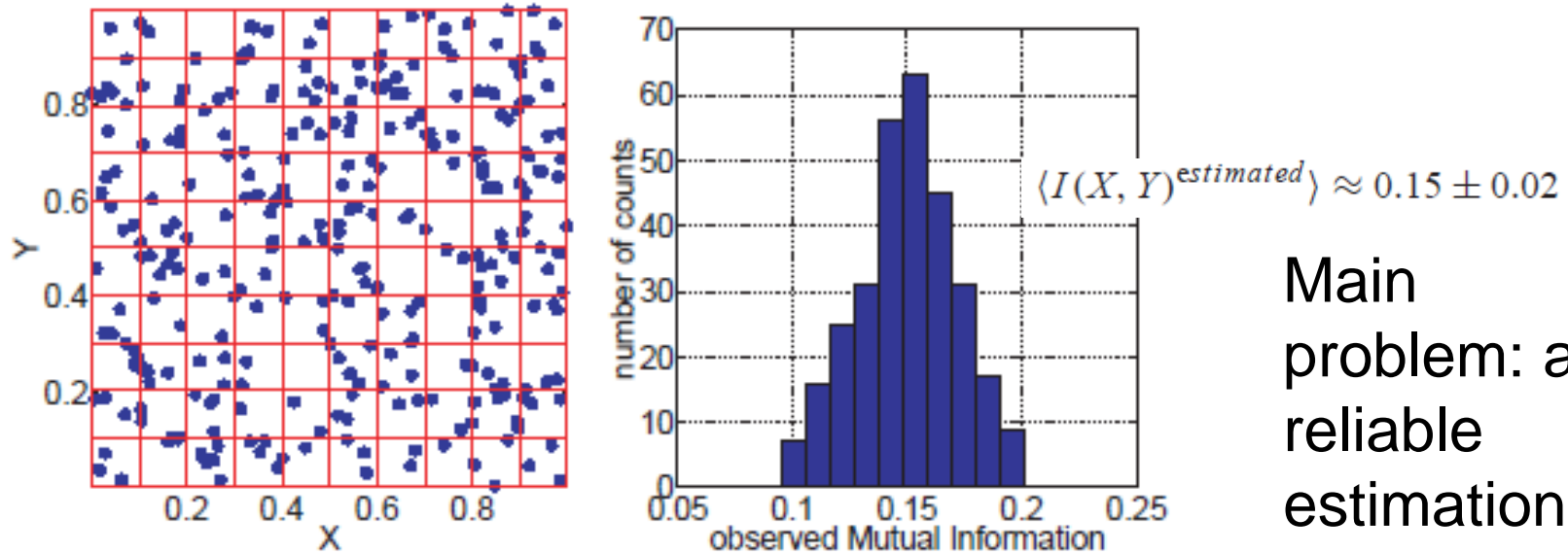


# Nonlinear correlation measure based on information theory: the mutual Information

$$MI = \sum_{i \in x} \sum_{j \in y} p(x, y) \log \left( \frac{p(x, y)}{p(x)p(y)} \right)$$

- $MI(x, y) = MI(y, x)$
- $p(x, y) = p(x) p(y) \Rightarrow MI = 0$ , else  **$MI > 0$**
- $MI$  can also be computed with a lag-time.
- $MI$  can also be computed from symbolic probabilities (e.g., probabilities of ordinal patterns).

# MI values are systematically overestimated

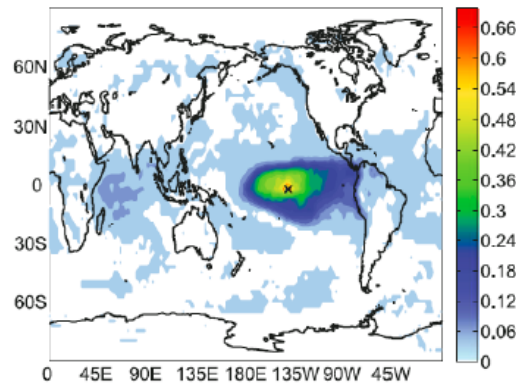


**Fig. 1.** Naive estimation of the mutual information for finite data. Left: The dataset consists of  $N = 300$  artificially generated independent and equidistributed random numbers. The probabilities are estimated using a histogram which divides each axis into  $M_x = M_y = 10$  bins. Right: The histogram of the estimated mutual information  $I(X, Y)$  obtained from 300 independent realizations.

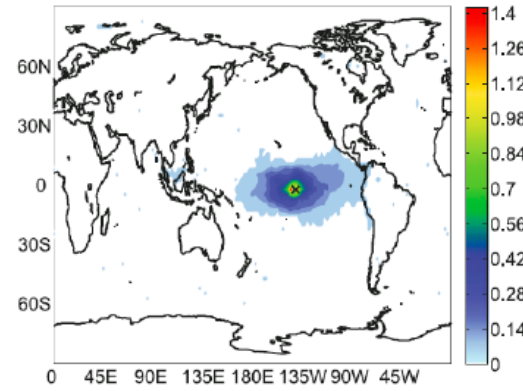
Main  
problem: a  
reliable  
estimation of  
MI requires a  
large amount  
of data

# Example: MI maps computed from SAT anomalies at a reference point located in El Niño, and all the other regions

Histograms



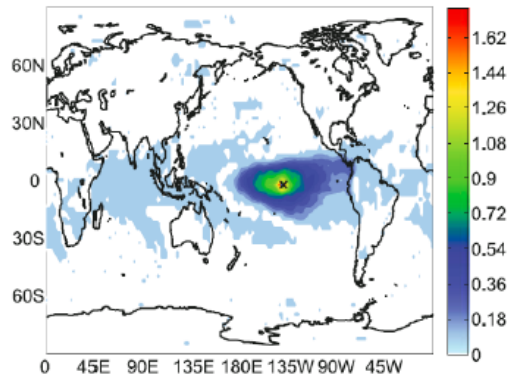
(a)



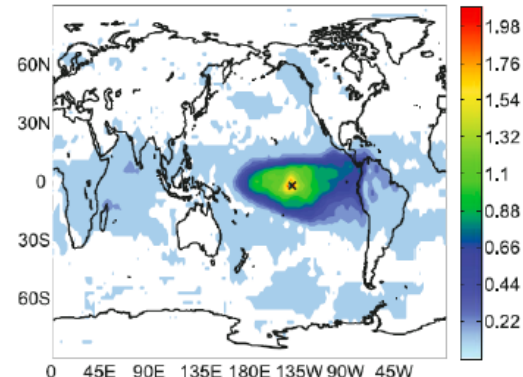
(b)

3 months  
ordinal  
patterns

Inter-  
annual  
ordinal  
patterns



(c)



(d)

3 years  
ordinal  
patterns

Ordinal analysis separates the times-scales of the interactions

# Direction of interaction?



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# Conditional mutual information (CMI) and transfer entropy (TE)

- CMI measures the amount of information shared between two time series  $i(t)$  and  $j(t)$ , given the effect of a third time series,  $k(t)$ , over  $j(t)$ .

$$M_I(i; j|k) = \sum_{m,n,l} p_{ijk}(m, n, l) \log \frac{p_k(l)p_{ijk}(m, n, l)}{p_{ik}(m, l)p_{jk}(n, l)}$$

- Transfer entropy = CMI with the third time series,  $k(t)$ , replaced by the *past* of  $i(t)$  or  $j(t)$ .

$$\text{TE}_{ij}(\tau) \equiv M_I(i; j|i_\tau) \quad \text{TE}_{ji}(\tau) \equiv M_I(j; i|j_\tau)$$

# Directionality index

- $\tau$ : *time-scale* of information transfer
- $DI$ : net direction of information transfer
- $DI_{ij} > 0 \rightarrow i$  drives  $j$ .

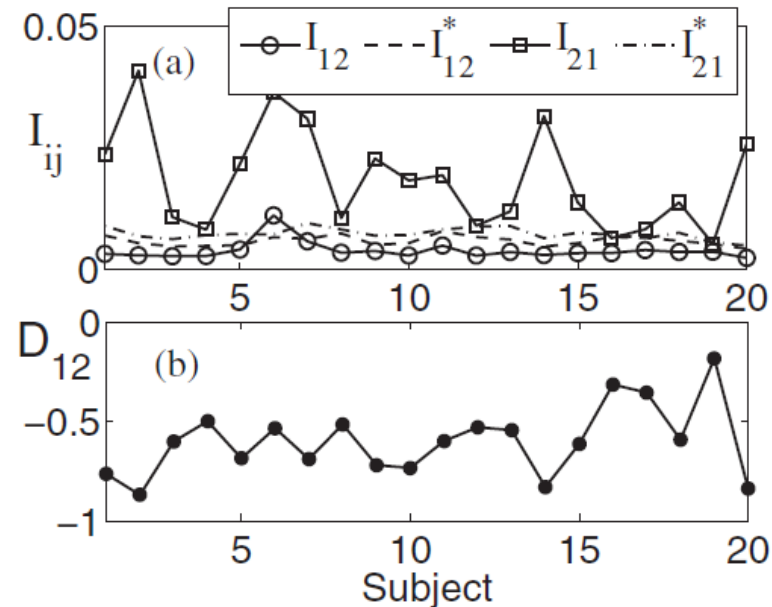
$$DI_{ij}(\tau) = \frac{TE_{ij}(\tau) - TE_{ji}(\tau)}{TE_{ij}(\tau) + TE_{ji}(\tau)}$$

Problem:  $x \rightarrow i$   
 $x \rightarrow j$      $i \leftrightarrow j$ ??

Application to **cardiorespiratory data** measured from 20 healthy subjects:  
(a) TEs (dashed lines: surrogate data)  
(b)  $D_{12}$  (1 = heart; 2 = respiration).

$D_{12} < 0 \rightarrow$  respiration drives cardiac activity.

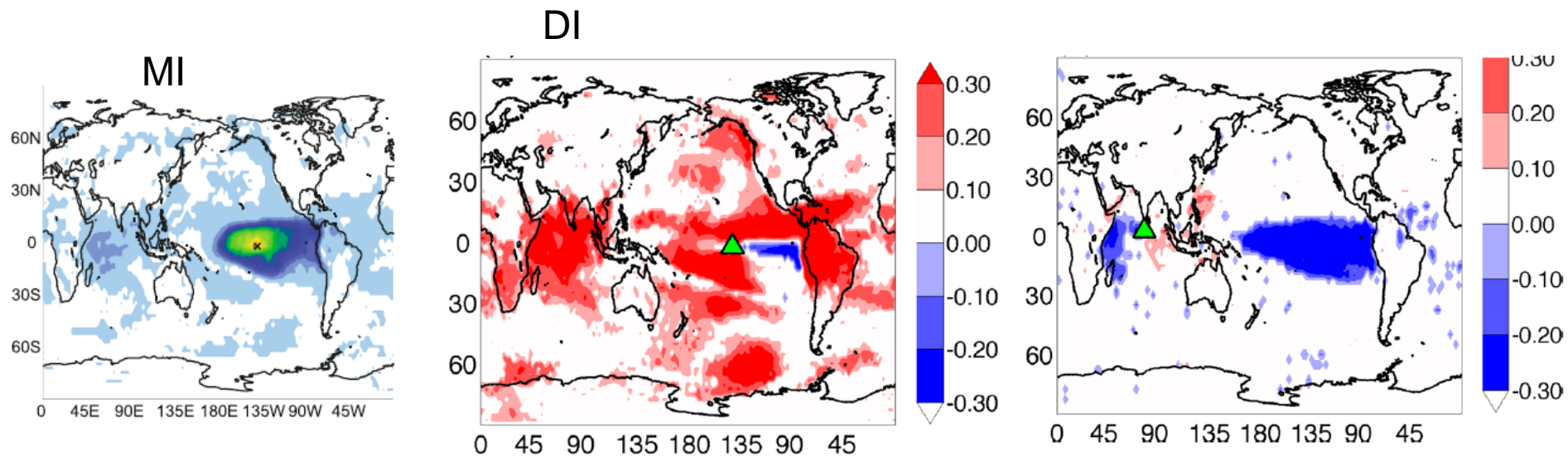
TEs were computed from ordinal probabilities and averaged over a short range of lags to decrease fluctuations.



# Application to climate data

DI computed from daily SAT anomalies, PDFs estimated from histograms of values.

MI and DI are both significant ( $>3\sigma$ , surrogates),  $\tau=30$  days.



[J. I. Deza, M. Barreiro, and C. Masoller, "Assessing the direction of climate interactions by means of complex networks and information theoretic tools", Chaos 25, 033105 \(2015\).](#)

# Causality?



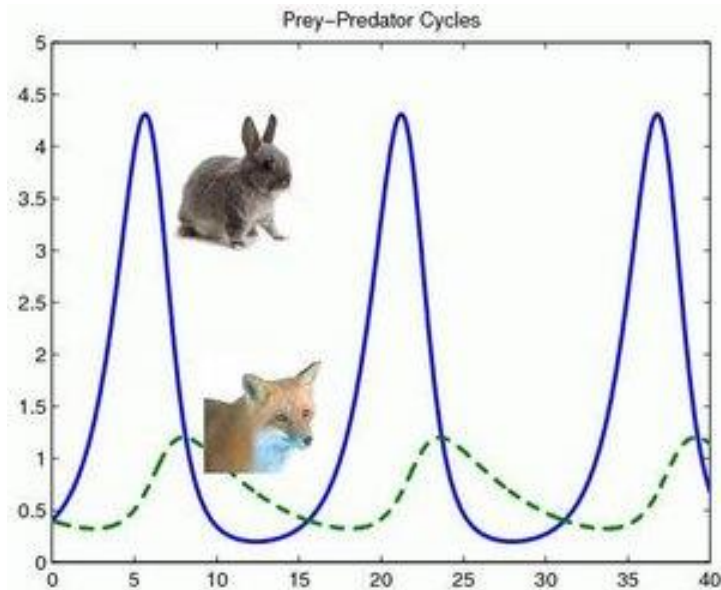
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# Main idea

- A time series  $X$  is Granger causal to a time series  $Y$  ( $X \rightarrow Y$ ) if the information given by  $X$  allows for a more precise prediction of  $Y$ .
- Example: in the predator - prey system, information about variations in the predator population can reveal properties of the prey population.



# Granger causality: how to detect $X \rightarrow Y$

- Model  $Y$  as a processes with memory forced by  $X$  with residual noise  $\varepsilon$

$$Y_t = \sum_{k=1}^D a_k Y_{t-k} + \sum_{k=1}^D b_k X_{t-k} + \varepsilon_t$$

- Test the hypothesis  $b \neq 0$  against the null hypothesis  $b=0$ :
  - Fit vectors  $a$  and  $b$  with a linear regression and compute the variance of the residual:  $\sigma_{\text{coupled}}^2$
  - Repeat with  $b=0$  and compute:  $\sigma_{\text{uncoupled}}^2$
  - Then calculate the Granger Causality Estimator

# Granger Causality Estimator

$$GCE = \frac{\sigma_{\text{uncoupled}}^2 - \sigma_{\text{coupled}}^2}{\sigma_{\text{uncoupled}}^2}$$

- If  $GCE > 0$  the information given by  $X$  allowed for a more precise prediction of  $Y$ .
- Problems:
  - how to select the dimension  $d$ ?
  - how to test the statistical significance of the GCE value?

# Summary

- Cross-correlation: detects linear interdependencies.
- Mutual information: detects nonlinear interdependencies.
- The MI computed from the probabilities of ordinal patterns allows to select the time-scale of the analysis.
- The directionality index detects the net direction of the information flow.
- Granger causality can “disentangle” mutual interactions.

## ■ Introduction

- Historical development: from dynamical systems to complex systems

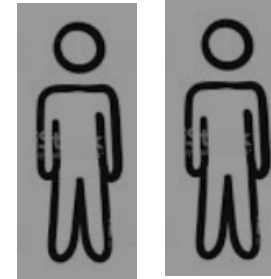
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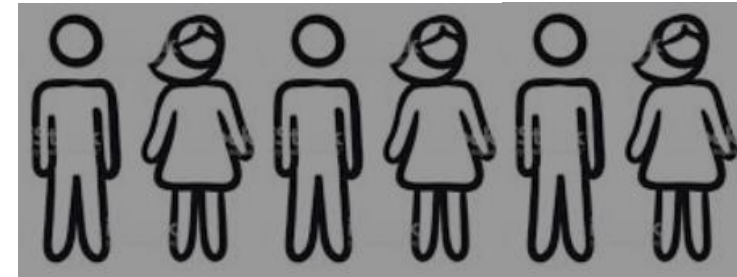
## ■ Bivariate analysis

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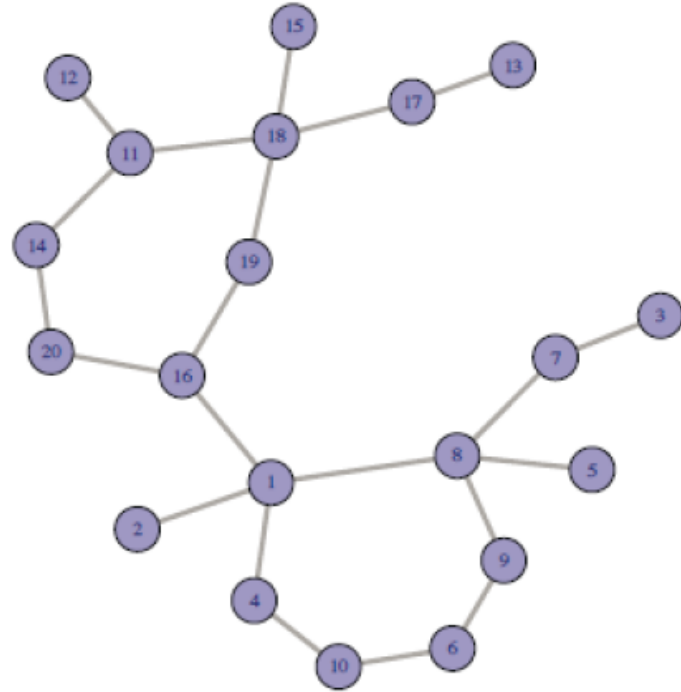
## ■ Multivariate analysis

- Many time series: complex networks.
- Network characterization and analysis.



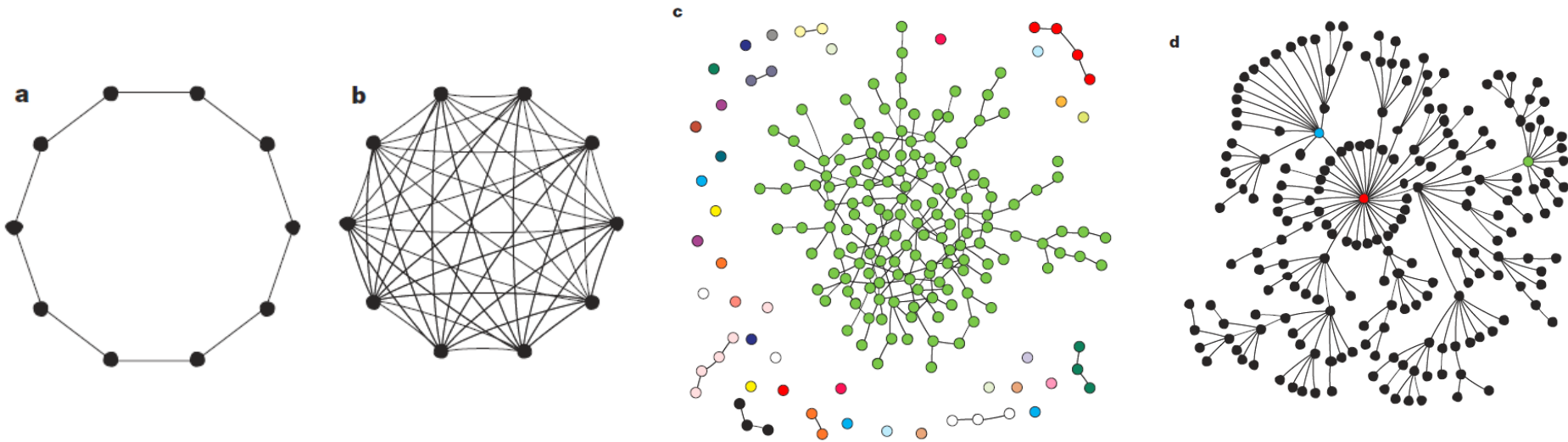
# What is a network?

- A graph: a set of “nodes” connected by a set of “links”.
- Nodes and links can be weighted or unweighted.
- Links can be directed or undirected.



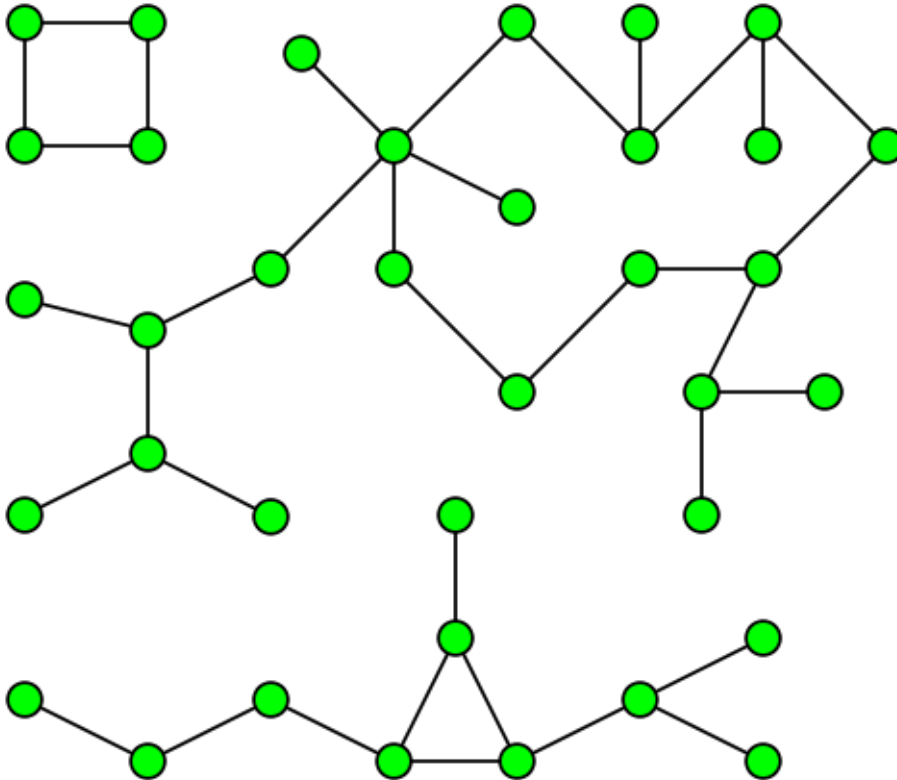
# Networks or graphs

The challenge in the context of time series analysis: to infer the underlying network structure from observed signals.



Source: Strogatz  
Nature 2001

# Connected components ("communities")

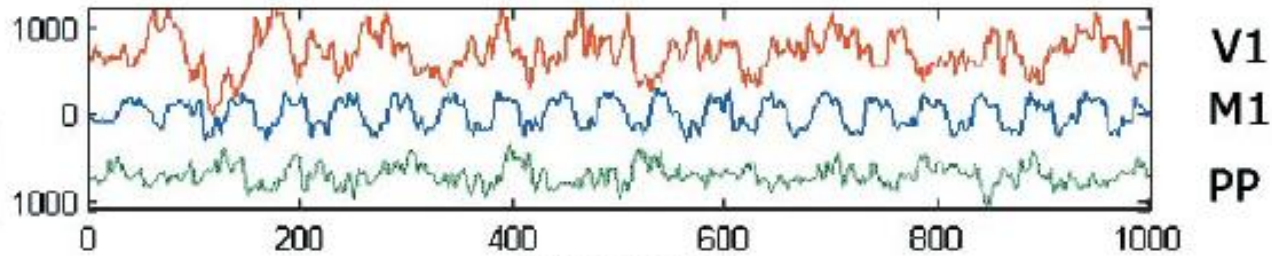


A graph with three connected components.  
Source: Wikipedia



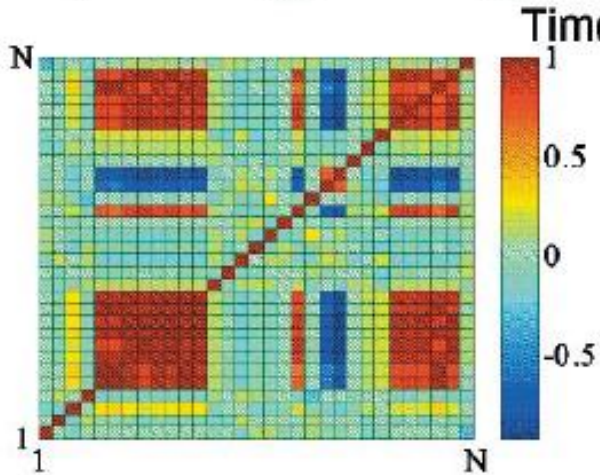
Using **bivariate** measures to  
infer interactions from data:  
“functional networks”

# Brain functional network

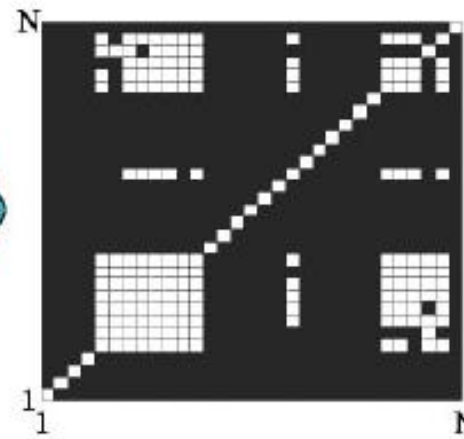


**Adjacency matrix**

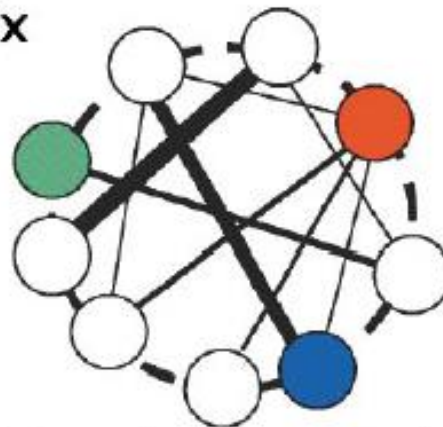
$$S_{ij} > Th \\ \Rightarrow A_{ij} = 1, \\ \text{else } A_{ij} = 0$$



Correlation Matrix



Thresholded Matrix

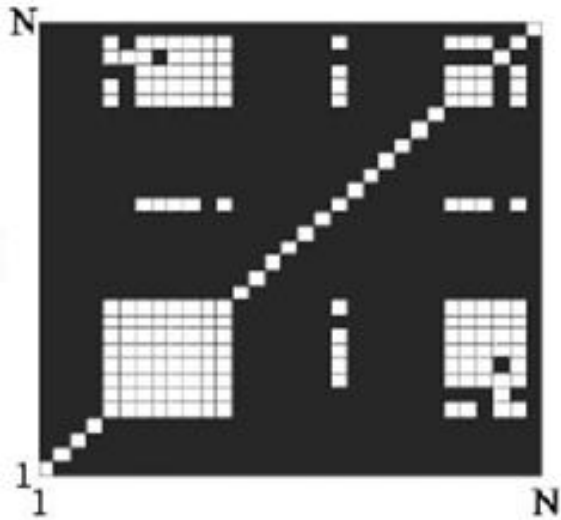


Network Extracted

*Eguiluz et al, PRL 2005*

# Graphical representation

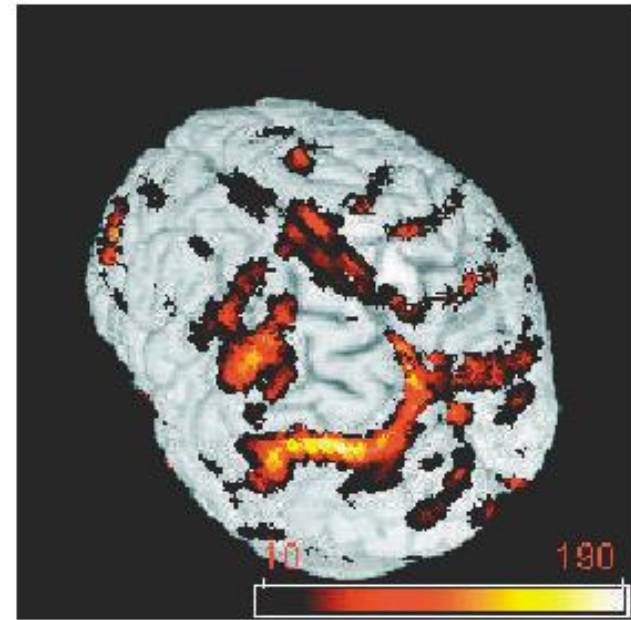
Adjacency matrix



Thresholded  
matrix = inferred  
("functional")  
network

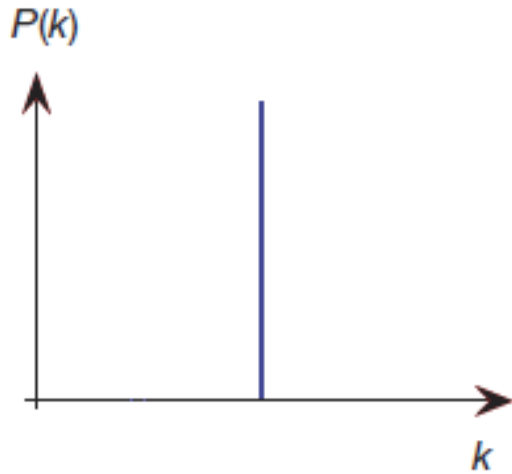
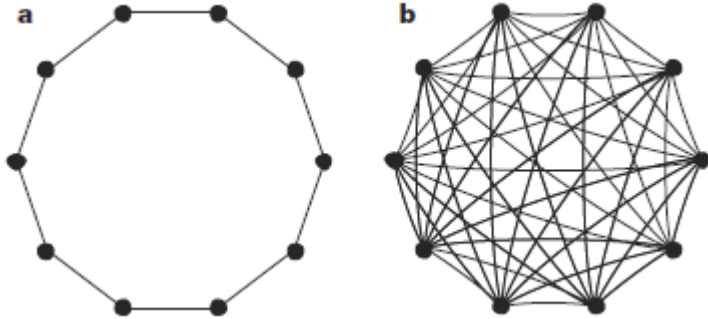
**Degree** of a node: number of links

$$k_i = \sum_j A_{ij}$$

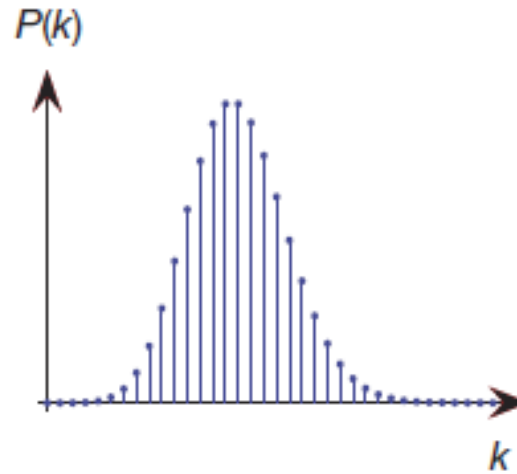
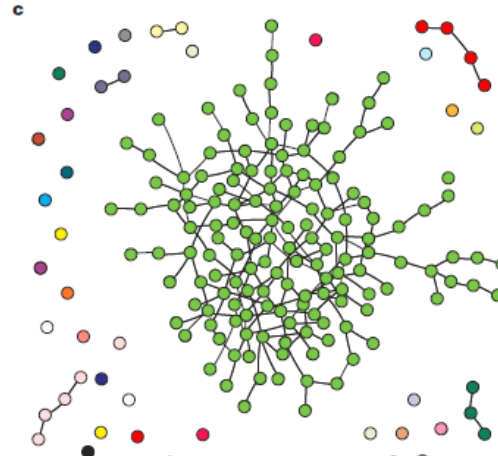


# The degree distribution: usual way to characterize a graph

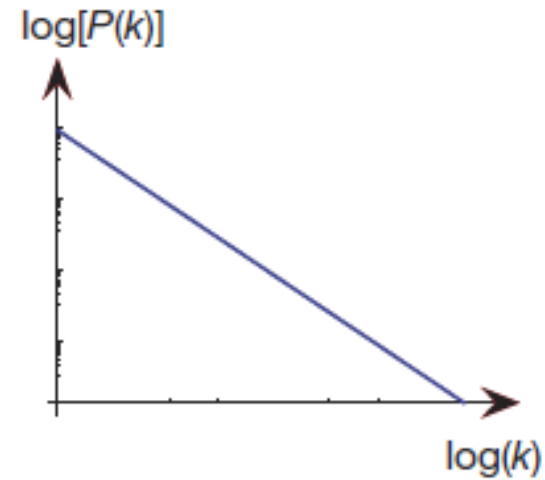
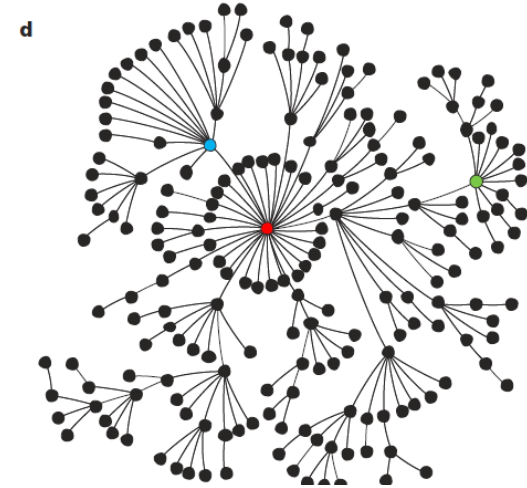
Regular



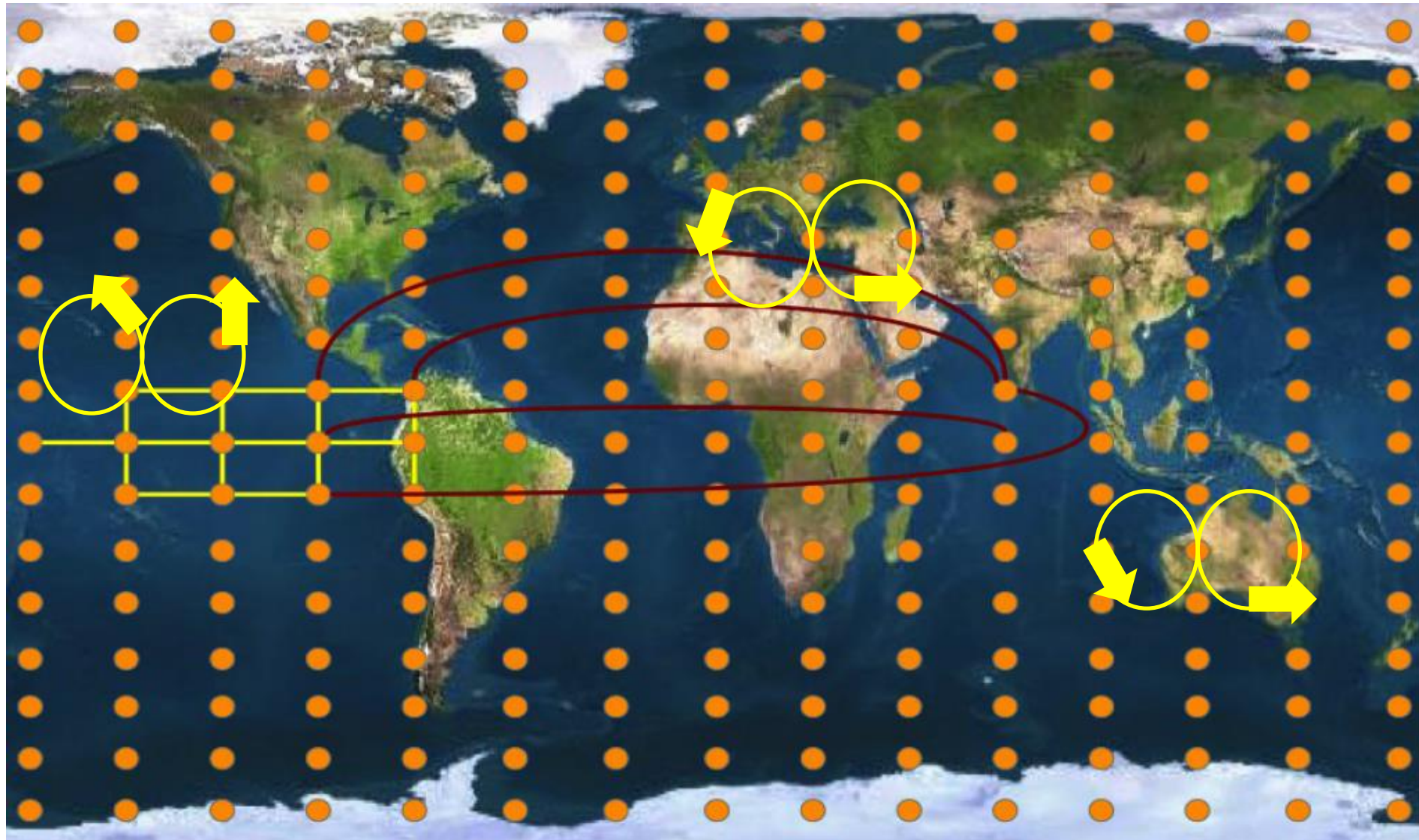
Random



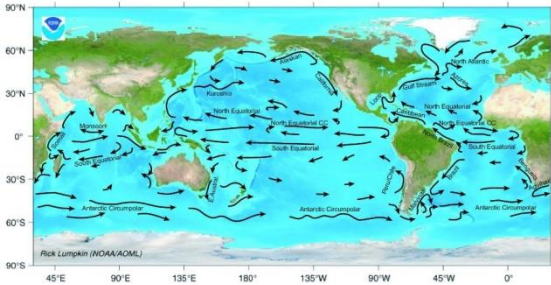
Scale-free



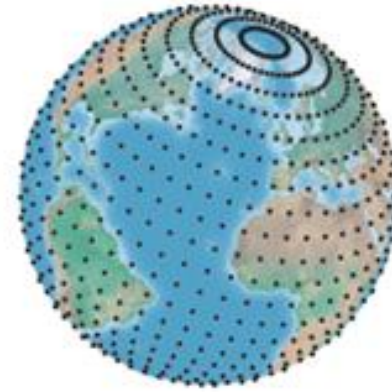
# The climate system as a set of “interacting oscillators”



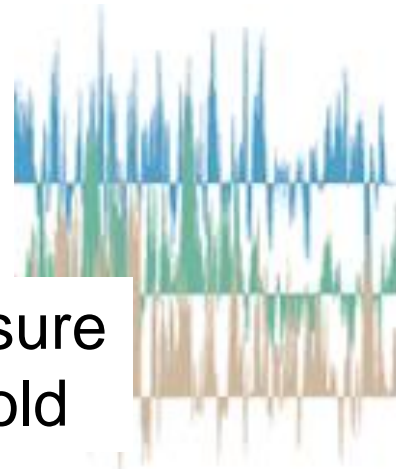
# Complex network representation of the climate system



Back to the climate system: interpretation (currents, winds, etc.)



More than 10000 nodes (with different sizes).



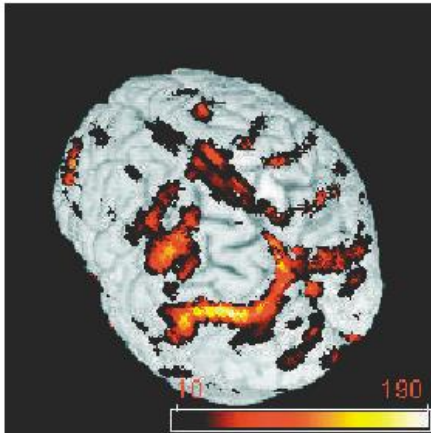
Daily resolution: more than 13000 data points in each TS



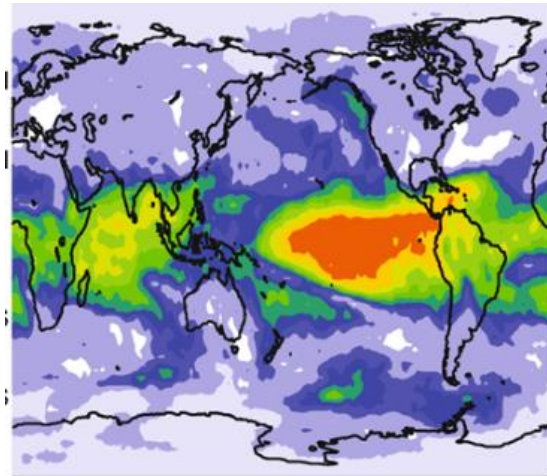
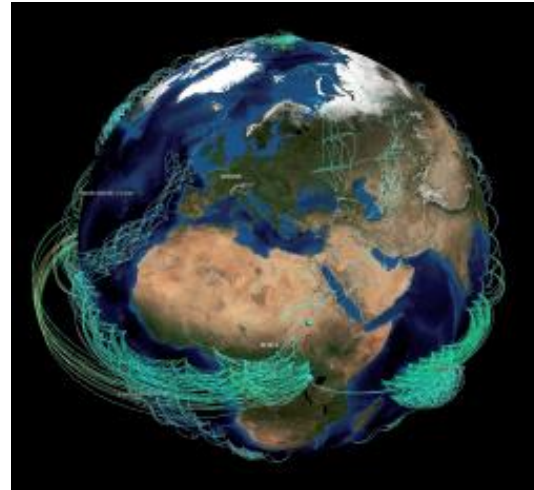
Sim. measure + threshold

Surface Air Temperature Anomalies (solar cycle removed)

## Brain network



## Climate network



Area weighted connectivity (AWC):  
weighted degree (nodes represent areas with different sizes)

$$AWC_i = \frac{\sum_j^N A_{ij} \cos(\lambda_j)}{\sum_j^N \cos(\lambda_j)}$$

## How to select the threshold ?

$$S_{ij} > Th \Rightarrow A_{ij} = 1, \\ \text{else } A_{ij} = 0$$

Three criteria are typically used:

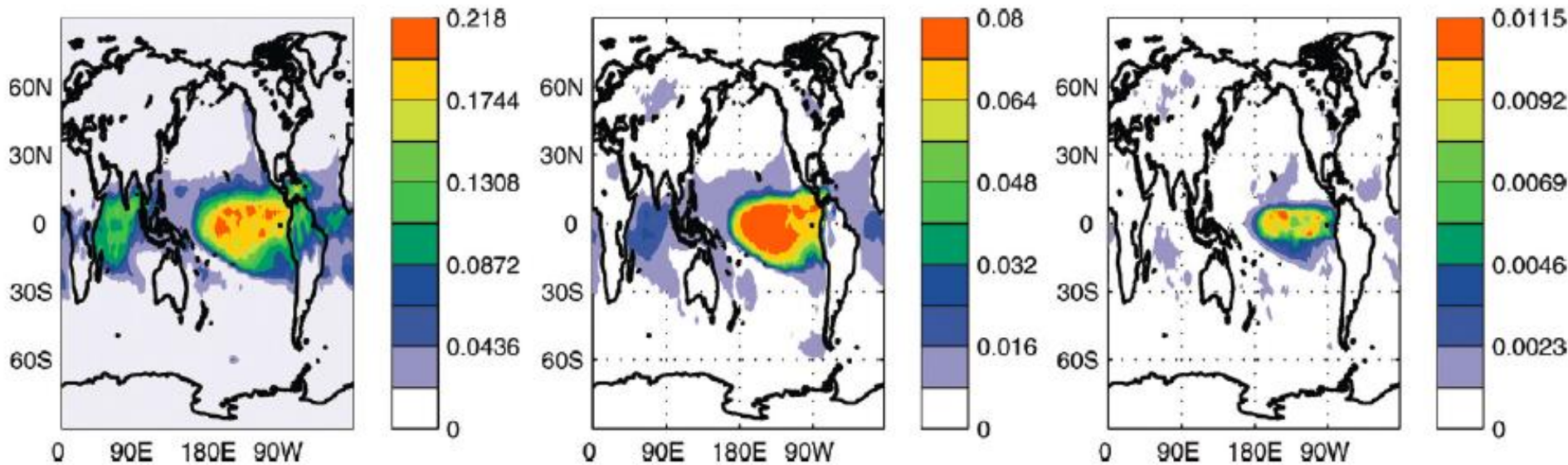
- A significance level is used (typically 5%) in order to omit connectivity values that can be expected by chance;
- We select an arbitrary value as threshold, such that it gives a certain pre-fixed number of links (or link density);
- We define the threshold as large as possible while guaranteeing that all nodes are connected (or a so-called “giant component” exists).



# How to select the threshold ?

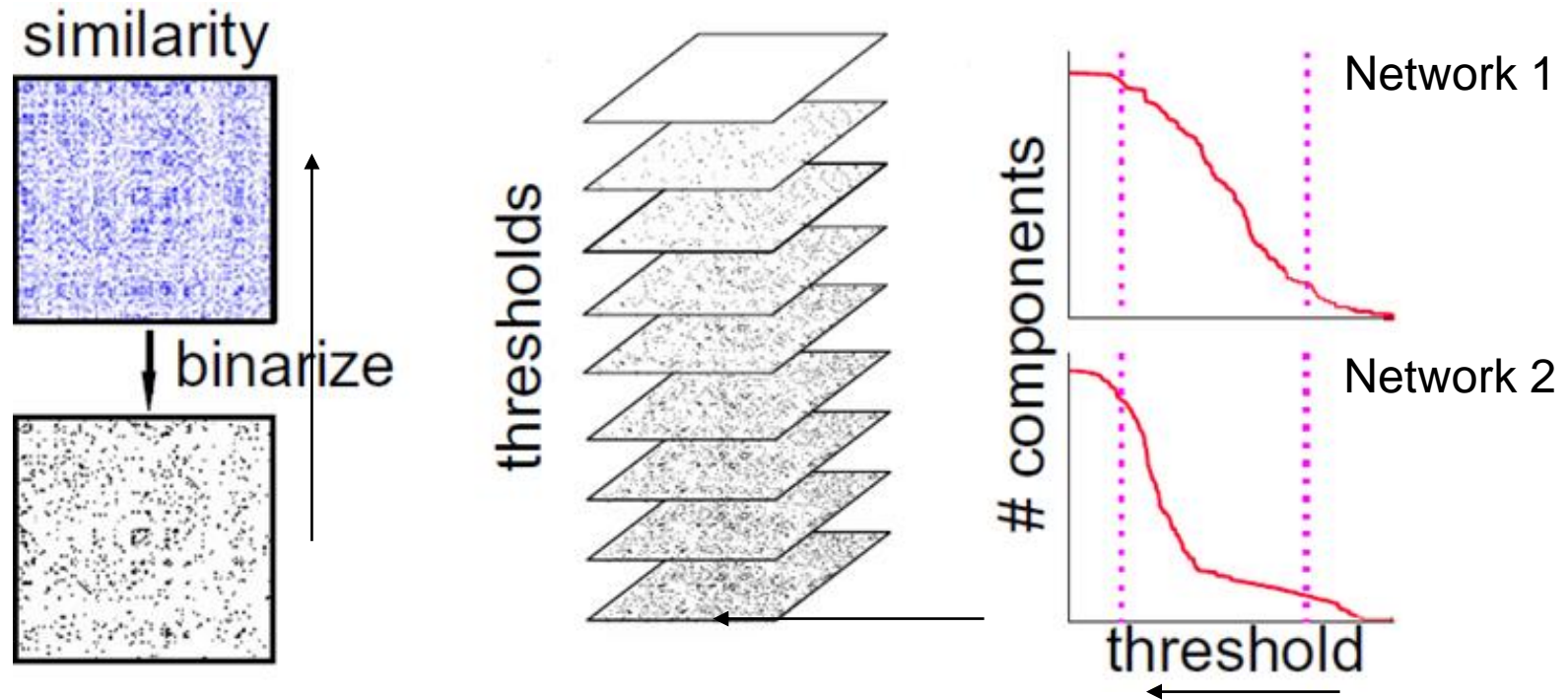
$$\text{If } S_{ij} > \text{Th} \Rightarrow A_{ij} = 1, \\ \text{else } A_{ij} = 0$$

Th  $\Rightarrow$



[M. Barreiro, et. al, Chaos 21, 013101 \(2011\)](#)

# Problems with thresholding



- The number of connected components as a function of threshold reveals different structures.
- But thresholding near the dotted lines indicates (inaccurately) that networks 1 and 2 have similar structures.

# **Network characterization**

## Definitions (for unweighted and undirected graphs)

- **Adjacency matrix:**  $A_{ij} = 1$  if  $i$  and  $j$  are connected, else  $A_{ij} = 0$ .

- **Degree** of a node  $k_i = \sum_j A_{ij}$

- **Clustering coefficient:** measures the fraction of a node's neighbors that are neighbors also among themselves

$$C_i = \frac{2R_i}{k_i(k_i - 1)} = \frac{1}{k_i(k_i - 1)} \sum_{j=1}^N \sum_{l=1}^N A_{ij} A_{jl} A_{li}$$

$R_i$  is the number of connected pairs in the set of neighbors of node  $i$

- **Assortativity:** measures the tendency of a node with high/low degree to be connected to other nodes with high/low degree

$$a_i \equiv \frac{1}{k_i} \sum_{j=1}^N A_{ij} k_j$$

# How to characterize the degree distribution?

- **Mean** (expected value of  $X$ ):  $\mu = E[X] = \sum_{i=1}^k x_i p_i = x_1 p_1 + x_2 p_2 + \dots + x_k p_k$

- **Variance**:  $\sigma^2 = \text{Var}(X) = E[(X - \mu)^2]$

- **Skewness**: “measures” the asymmetry of the distribution

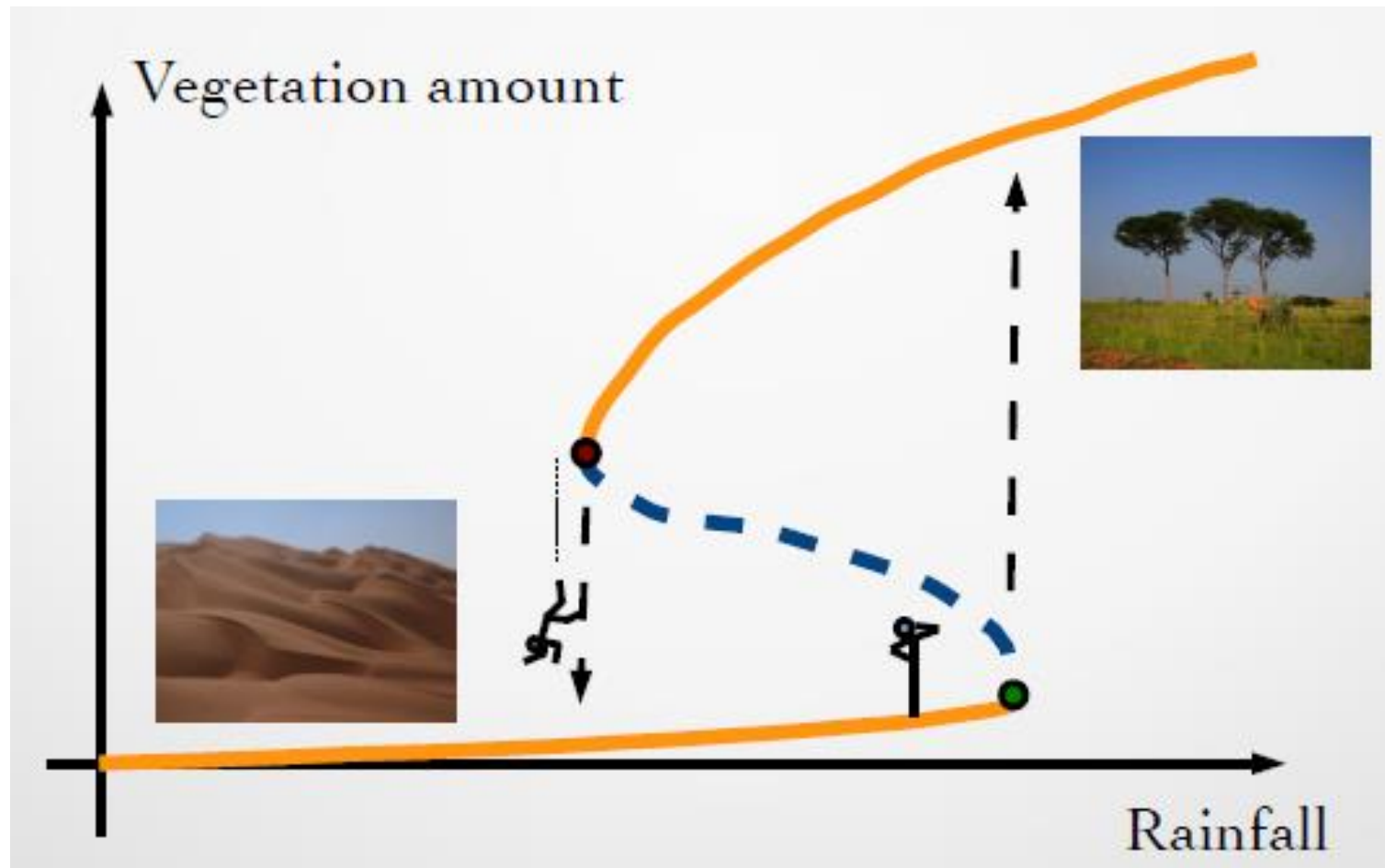
$$Z = \frac{X - \mu}{\sigma} \quad S = E[Z^3]$$

- **Kurtosis**: measures the “tailedness” of the distribution. For a normal distribution  $K=3$ .

$$K = E[Z^4]$$



**Example of application:  
desertification transition**



Our goal: to develop reliable early-warning indicators

Can we use “correlation networks” to detect the approach to a tipping point?

# Model

$$\frac{\partial w}{\partial t} = R - \frac{w}{\tau_w} - \Lambda w B + D \nabla^2 w + \sigma_w w_0 \xi^w(t),$$

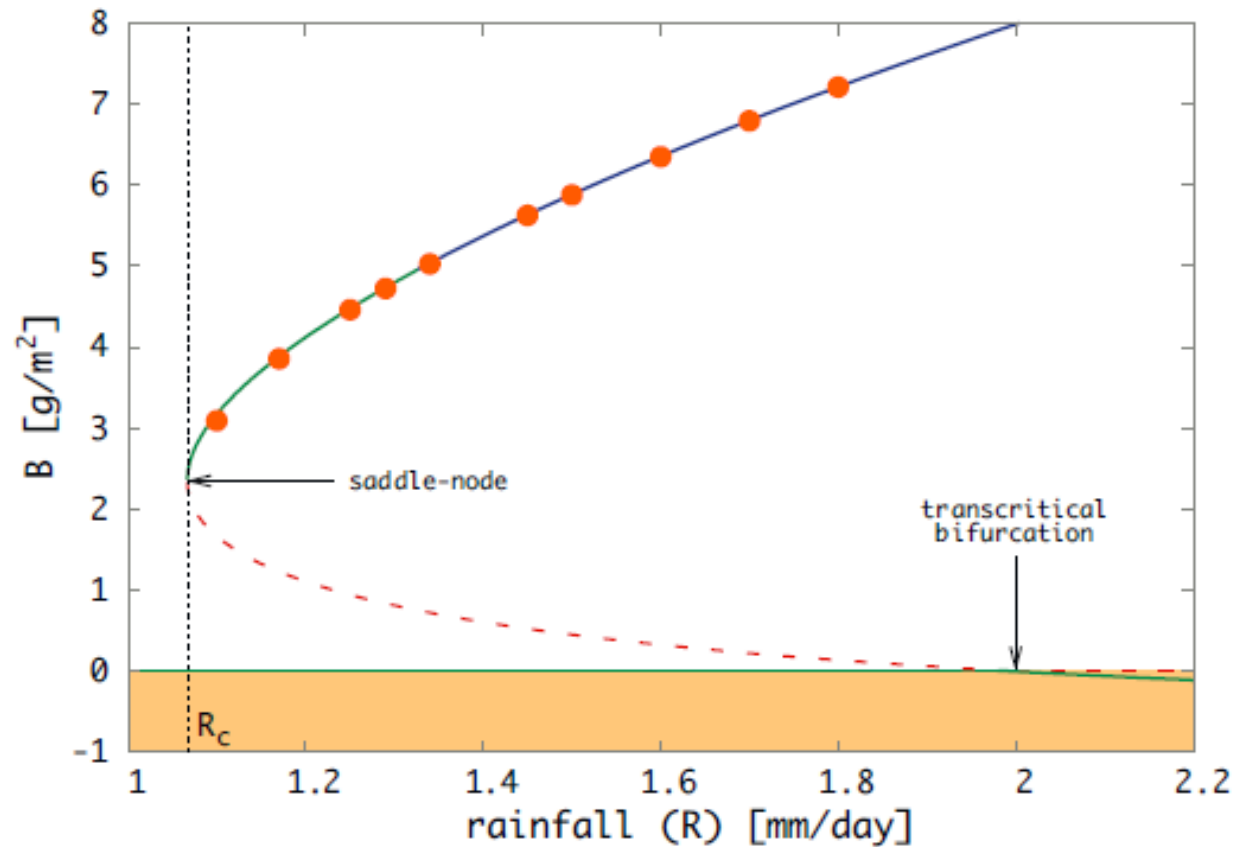
$$\frac{\partial B}{\partial t} = \rho B \left( \frac{w}{w_0} - \frac{B}{B_c} \right) - \mu \frac{B}{B + B_0} + D \nabla^2 B + \sigma_B B_0 \xi^B(t)$$

- $w$  (in mm) is the soil water amount
- $B$  (in g/m<sup>2</sup>) is the vegetation biomass
- Uncorrelated Gaussian white noise
- $R$  (rainfall) is the bifurcation parameter

*Shnerb et al. (2003), Guttal & Jayaprakash (2007), Dakos et al. (2011)*



# Saddle-node bifurcation

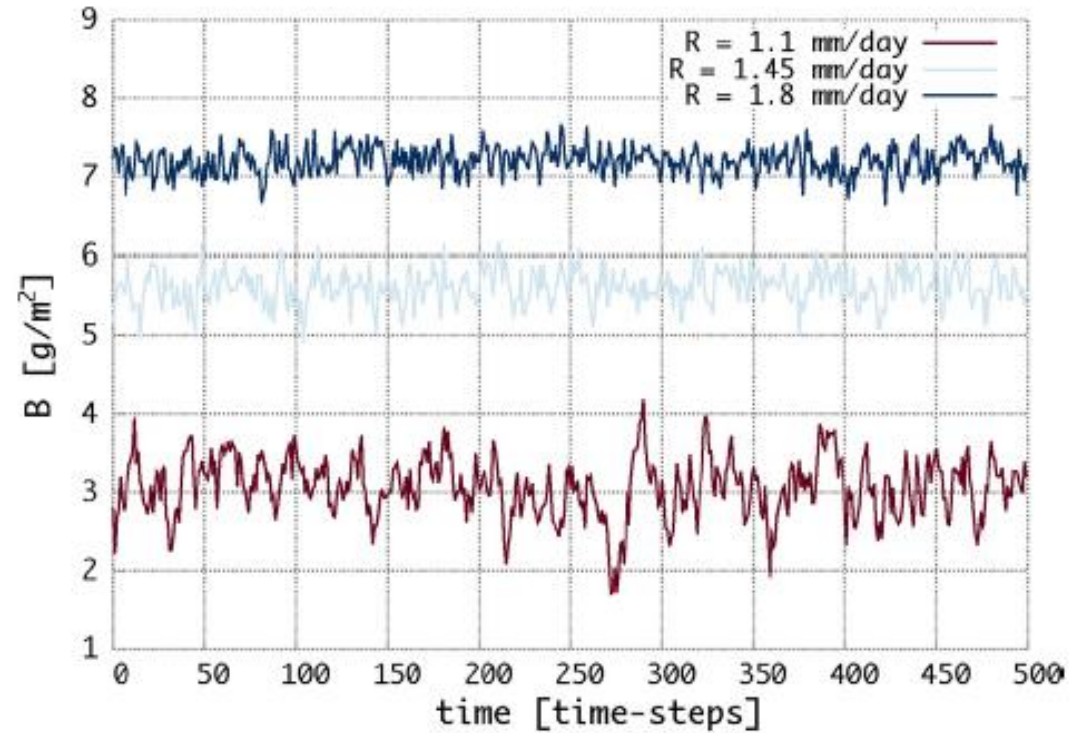
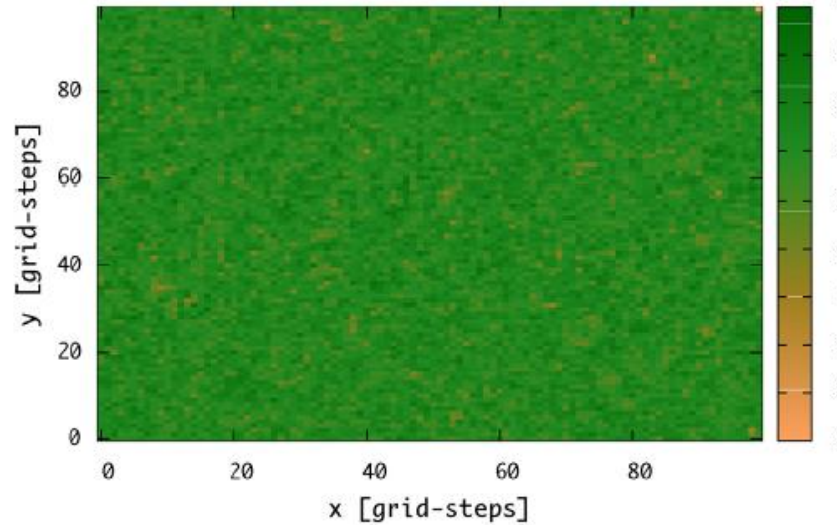


$R < R_c$ : only desert-like solution ( $B=0$ )

**$R_c = 1.067$  mm/day**

# Biomass time series

Biomass  $B$  when  $R=1.1$  mm/day



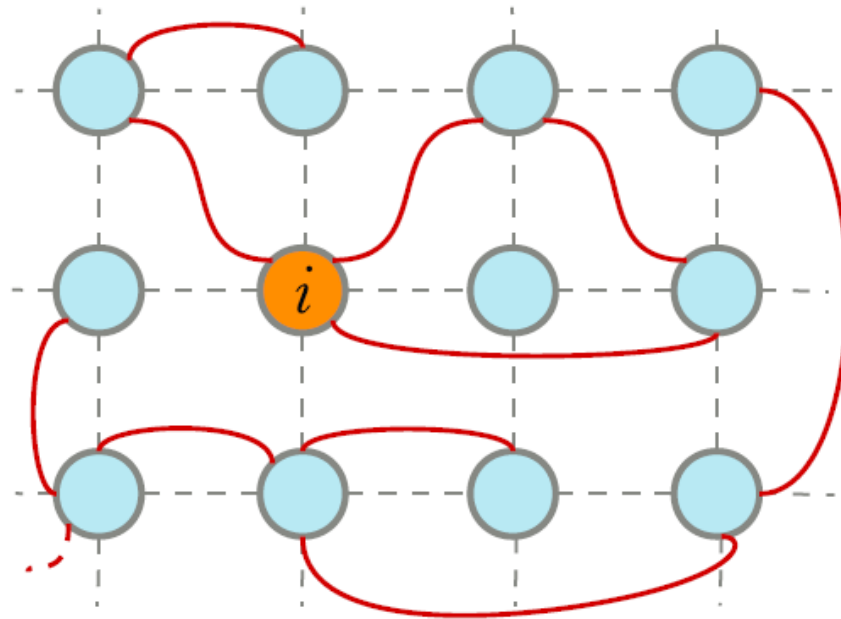
100 m x 100 m =  $10^4$  grid cells  
Simulation time 5 days in 500 time steps  
Periodic boundary conditions

# Correlation Network

$$S_{ij} > Th \Rightarrow A_{ij} = 1, \text{ else } A_{ij}=0$$

Statistical similarity  
measure:  
Pearson coef.=  
|zero-lag cross-  
correlation|

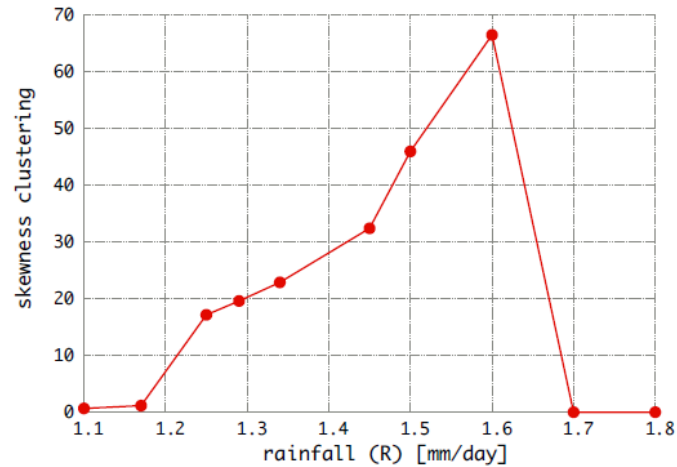
Threshold:  $Th=0.2$  keeps only  
significant correlations ( $p<0.05$ )



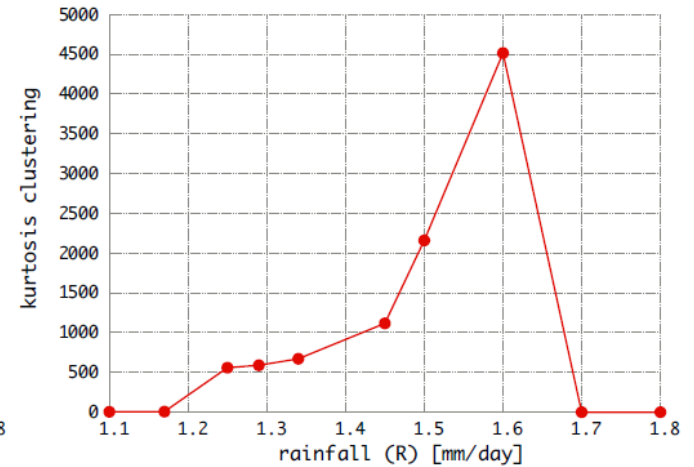
# “Gaussianization” of the distributions of $a_i$ & $c_i$ as the tipping point is approached

clustering

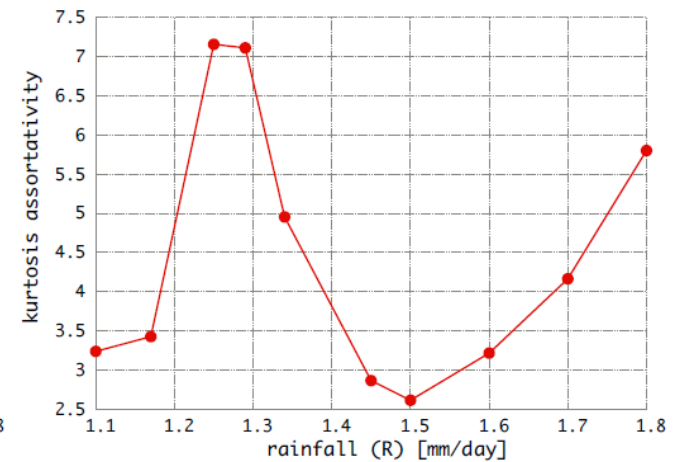
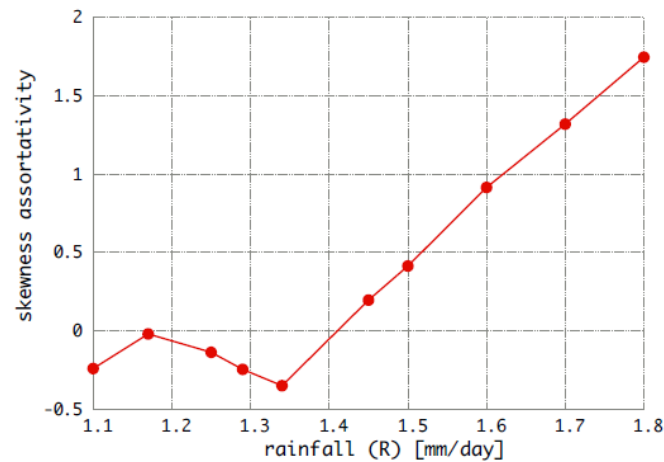
## skewness



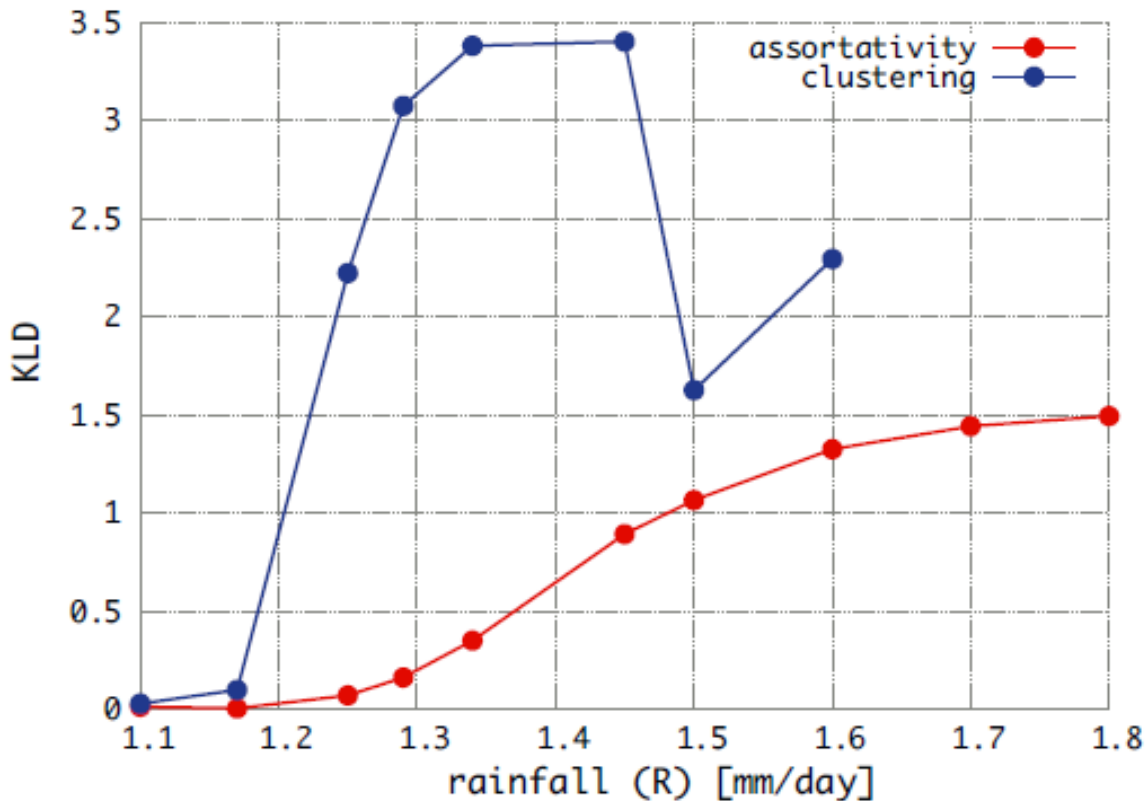
## kurtosis



assortativity



# The “Gaussianisation” is quantified by the Kullback distance to a Gaussian (Z) distribution



$$\text{KLD} \equiv \int_{-\infty}^{\infty} \ln \left( \frac{P(x)}{Z(x)} \right) P(x) dx.$$

- Open issue: the “Gaussianisation” might be a model-specific feature.
- How to precisely quantify changes of the network?
- We need a distance to compare graphs.

**How to “infer” interactions  
from observed data?**

# A classification problem

$$S_{ij} > Th \Rightarrow A_{ij} = 1, \text{ else } A_{ij}=0$$

- How to select the threshold?
- In “spatially embedded networks”, nearby nodes have the strongest links.
- How to keep **weak-but-significant** links?
- There are many **statistical similarity measures** to infer interactions from observations, i.e., to classify:
  - the interaction exists (is significant)
  - the interaction does not exist (or is not significant)

**Goal: use a system with known connectivity to test the performance of statistical similarity measures**

*Observed time series in nodes  $i$  and  $j$ :  $a_i(t)$ ,  $a_j(t)$ ,  $t=1, \dots, T$   
(normalized  $\mu=0$ ,  $\sigma=1$ )*

Lagged |cross correlation|: 
$$CC_{ij}(\tau) = \frac{1}{T - \tau_{\max}} \left| \sum_{t=0}^{T-\tau_{\max}} a_i(t) a_j(t + \tau) \right|$$

Statistical Similarity Measure:

$$\begin{aligned} S_{ij} &= \max | CC_{ij}(\tau) | \\ &= | CC_{ij}(\tau_{ij}) | \quad \tau_{ij} \text{ in } [0, \tau_{\max}] \end{aligned}$$

We compare with the Mutual Information, computed from probabilities of “raw” values and from ordinal probabilities

[G. Tirabassi et al., “Inferring the connectivity of coupled oscillators from time-series statistical similarity analysis”, Sci. Rep. 5 10829 \(2015\).](#)



# Kuramoto oscillators in a random network

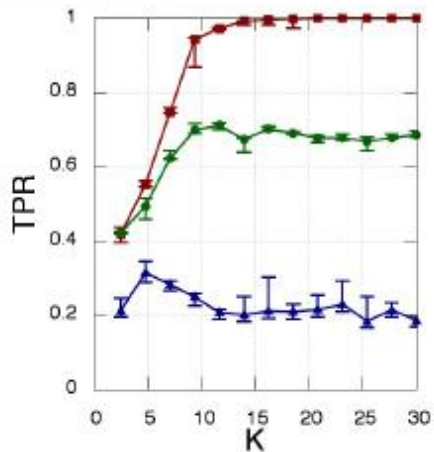
$$d\theta_i = \omega_i dt + \frac{K}{N} \sum_{j=1}^N A_{ij} \sin(\theta_j - \theta_i) dt + D dW_t^i$$

$A_{ij}$  is a symmetric random matrix;  
 $N=12$  time-series, each with  $10^4$  data points.

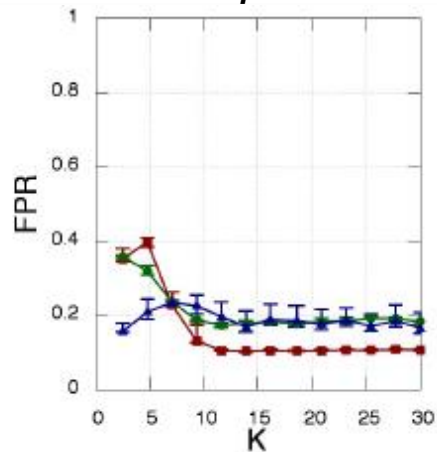
## Phases ( $\theta$ )

CC MI MIOP

### True positives

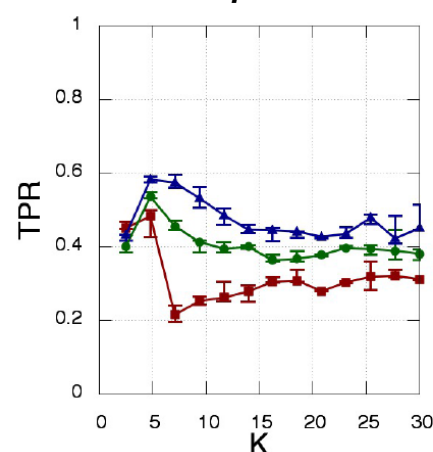


### False positives

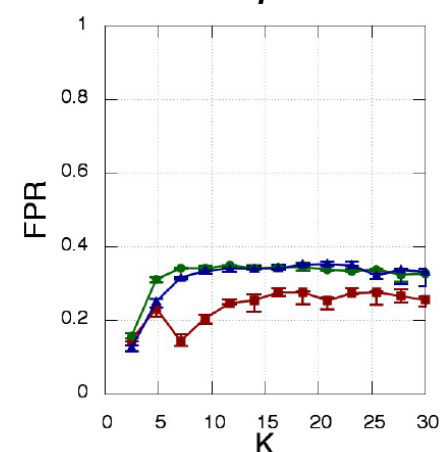


## “Observable” $Y=\sin(\theta)$

### True positives



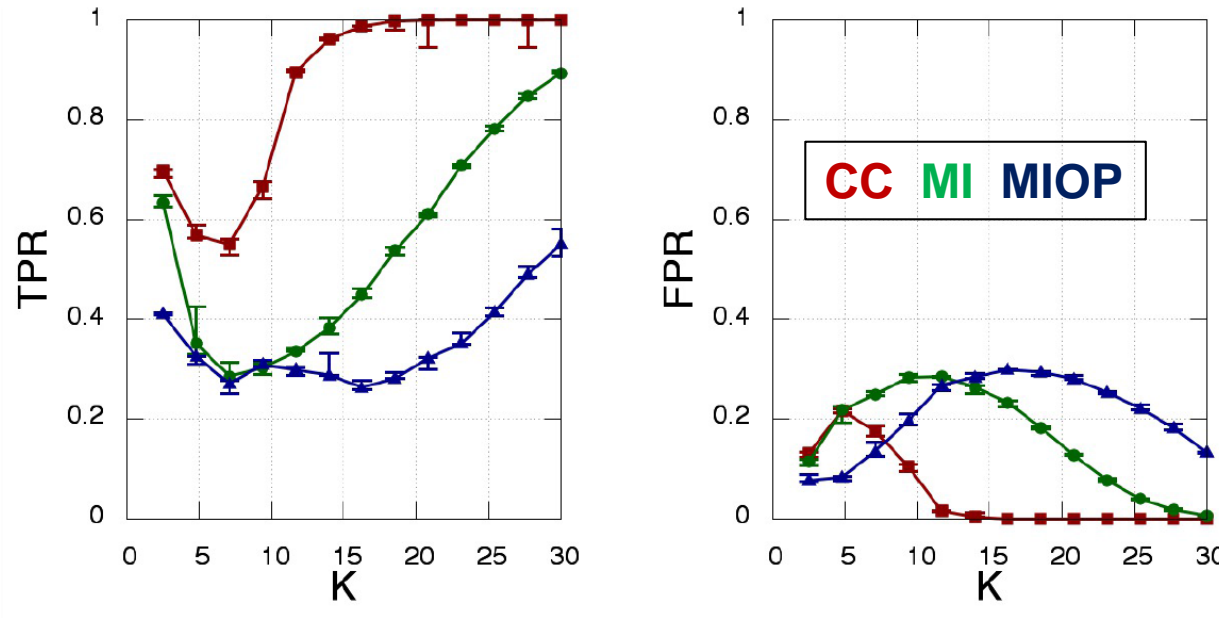
### False positives



Results of a 100 simulations with different oscillators' frequencies, random matrices, noise realizations and initial conditions.

For each  $K$ , the threshold was varied to obtain optimal reconstruction.

# Instantaneous frequencies ( $d\theta/dt$ )



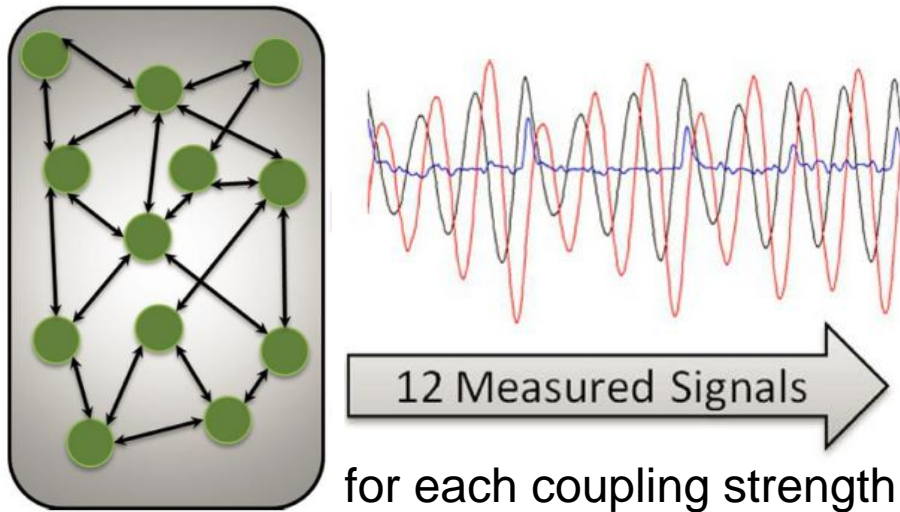
Perfect network inference is possible!

BUT

- the number of oscillators is small (12),
- the coupling is symmetric ( $\Rightarrow$  only 66 possible links) and
- the data sets are long ( $10^4$  points)

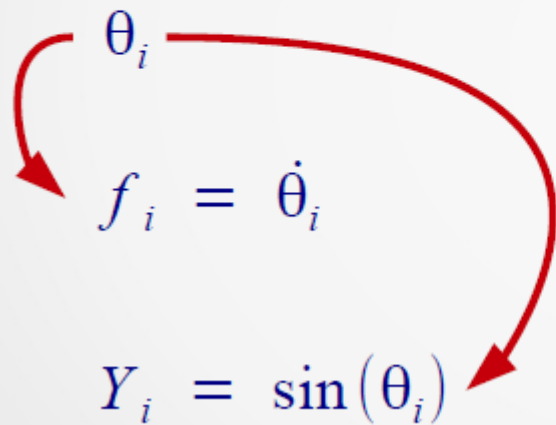
[G. Tirabassi et al, Sci. Rep. 5 10829 \(2015\)](#)

We also analyzed experimental data recorded from 12 chaotic Rössler electronic oscillators (symmetric and random coupling)

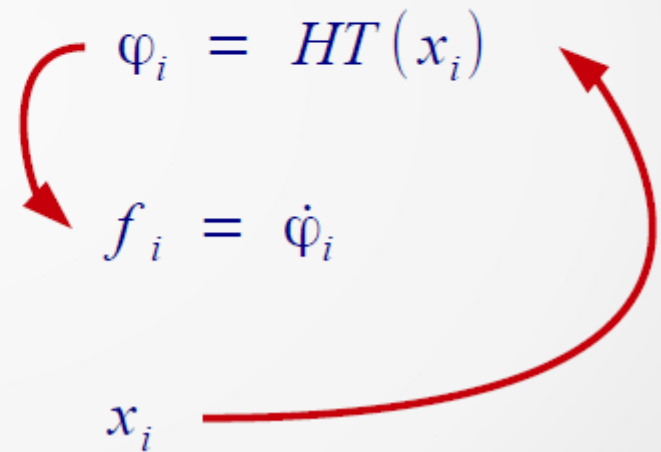


The Hilbert Transform was used to obtain phases from experimental data

- Kuramoto Oscillators' Network

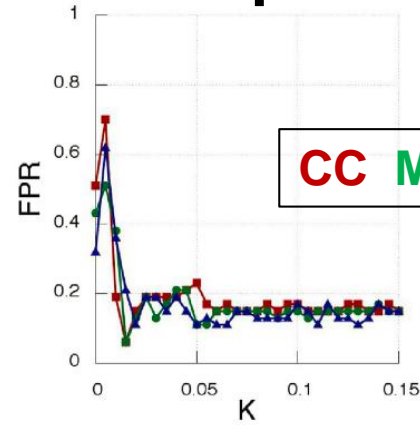
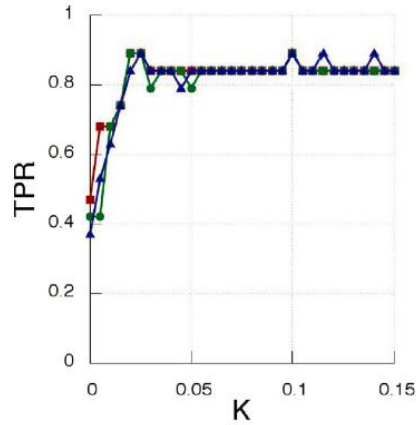


- Rössler Oscillators' Network

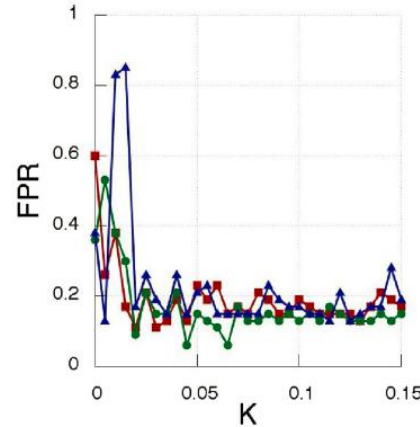
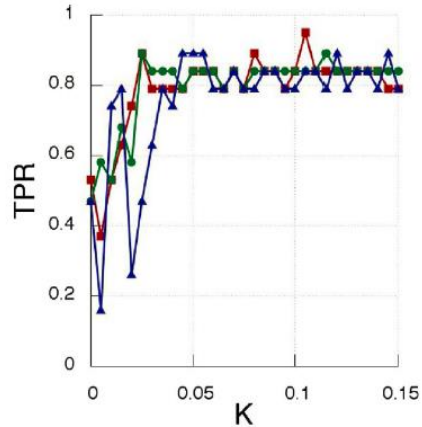


# Results obtained with experimental data

Observed variable (x)



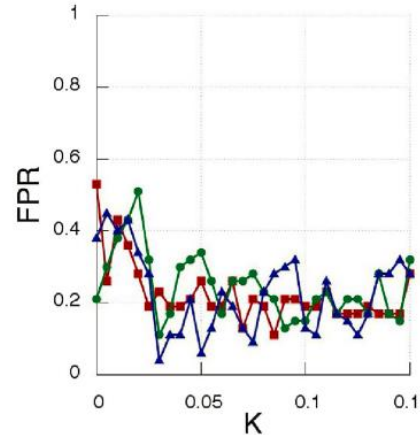
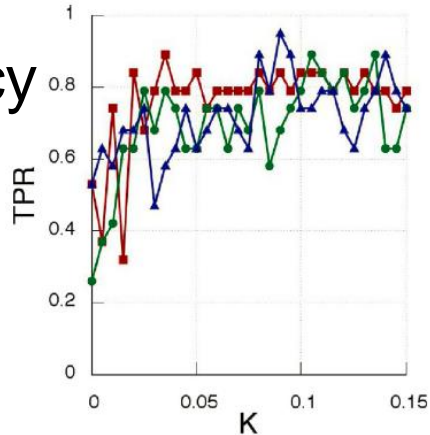
Hilbert phase



– No perfect reconstruction

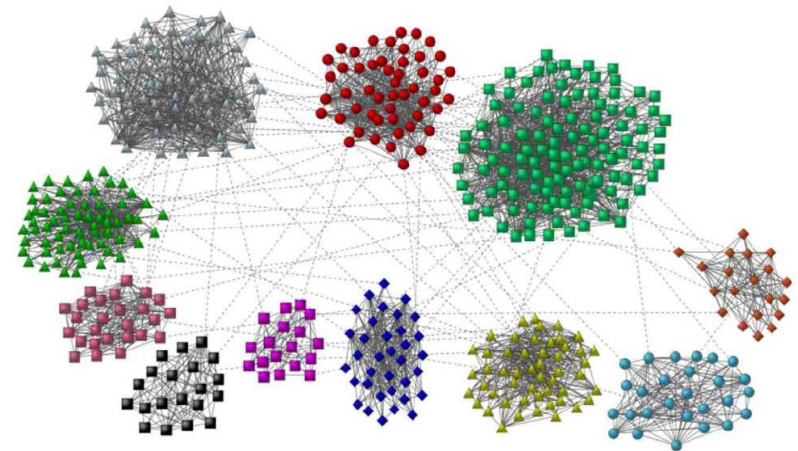
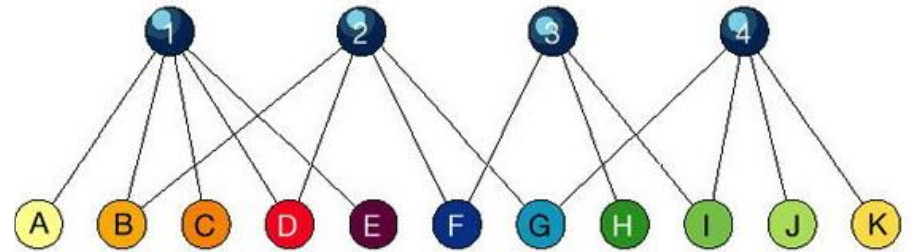
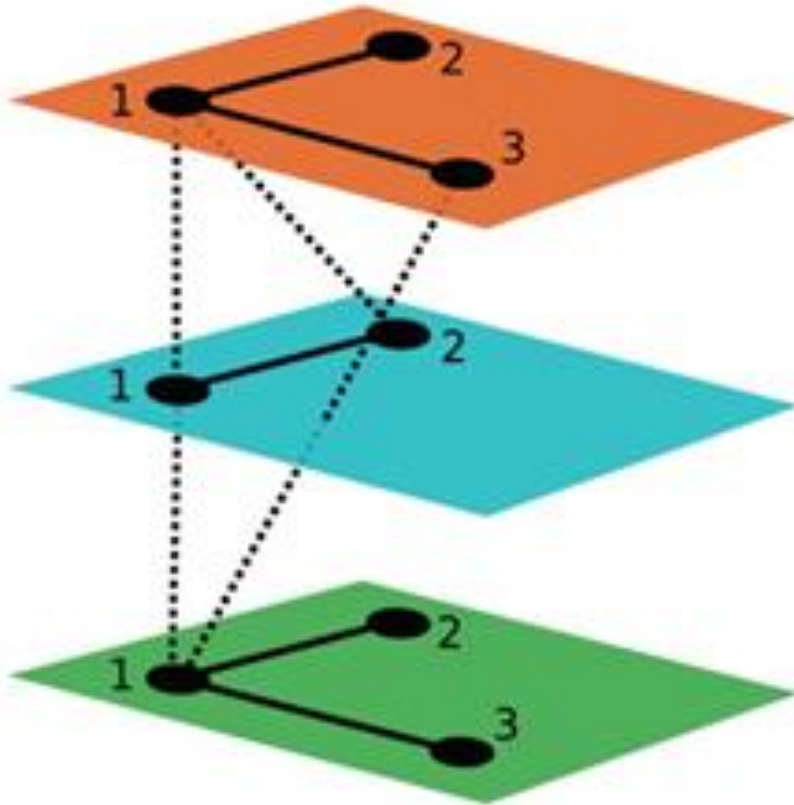
– No important difference among the 3 methods & 3 variables

Hilbert frequency



# **Generalizations of complex network analysis**

# Network structures: Multilayer, multiplex, bipartite, networks of networks and many others

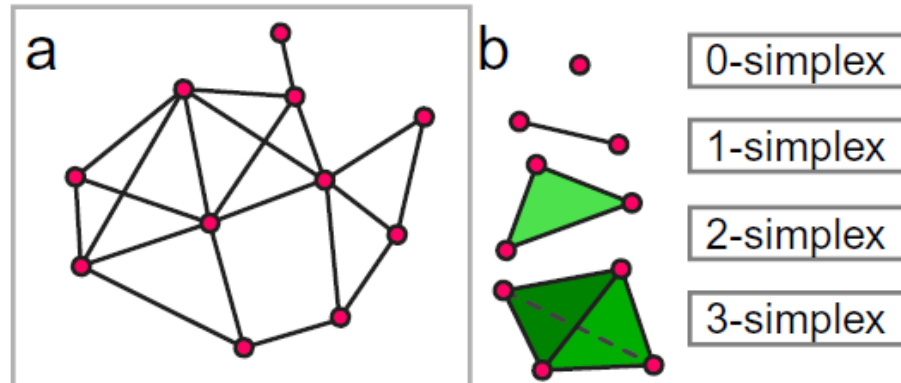


# **Limitations of complex network analysis**

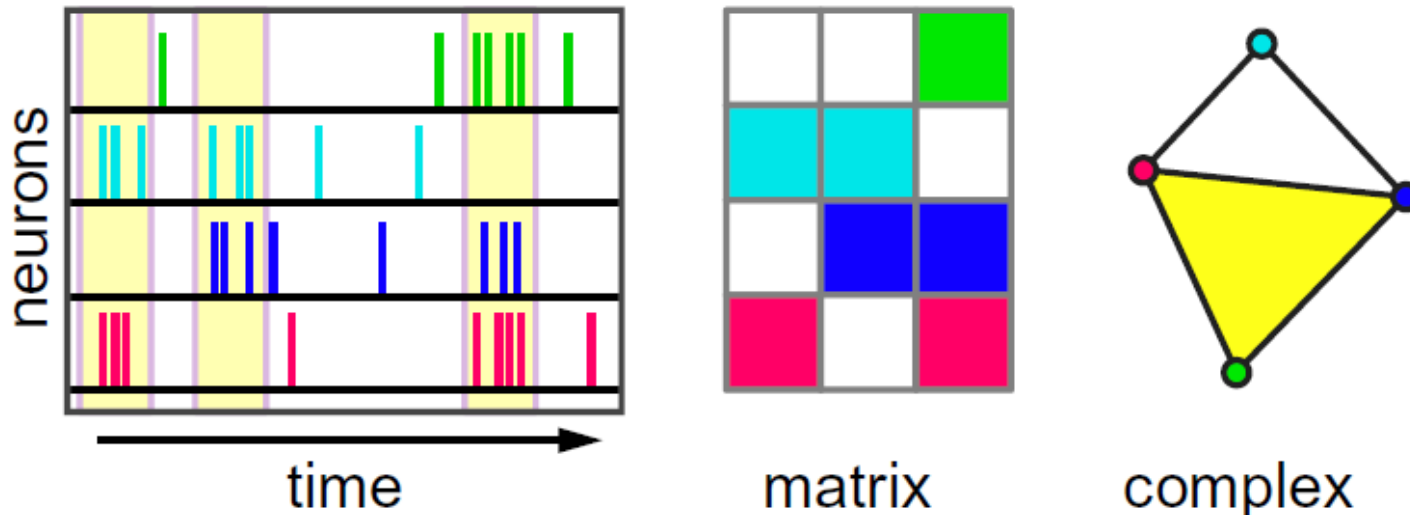


# Interactions are not limited to pairs of elements

- Links represent interactions between pairs of nodes.
- **Simplicial complexes** represent interactions among several nodes.



## Example



# Summary

- Multivariate analysis uncovers inter-relationships in datasets
- Different similarity measures are available for inferring the connectivity of a complex system from observations.
- Different measures can uncover different properties.
- Thresholding, hidden variables, hidden “nodes” can difficult or make impossible the inference of the network structure.
- Different sets of “communities” can be uncovered depending on the property that is analyzed.
- Network science: many applications and challenges!

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